

Causes of the long-term variability of southwestern South America precipitation in the IPSL-CM6A-LR model

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Julian Villamayor, Myriam Khodri, Ricardo Villalba, Valérie Daux. Causes of the long-term variability of southwestern South America precipitation in the IPSL-CM6A-LR model. Climate Dynamics, 2021, 10.1007/s00382-021-05811-y. insu-03230914

HAL Id: insu-03230914 https://insu.hal.science/insu-03230914v1

Submitted on 20 May 2021

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6 Received: date / Accepted: date

Abstract Southwestern South America (SWSA) has undergone frequent and per-7 sistent droughts in recent decades with severe impacts on water resources, and 8 consequently, on socio-economic activities at a sub-continental scale. The local 9 drying trend in this region has been associated with the expansion of the sub-10 tropical drylands over the last decades. It has been shown that SWSA precipi-11 tation is linked to large-scale dynamics modulated by internal climate variability 12 and external forcing. This work aims at unravelling the causes of this long-term 13 trend toward dryness in the context of the emerging climate change relying on a 14 large set simulations of the state-of-the-art IPSL-CM6A-LR climate model from 15 the 6th phase of the Coupled Model Intercomparison Project. Our results iden-16 tify the leading role of dynamical changes induced by external forcings, over the 17 local thermodynamical effects and teleconnections with internal global modes of 18 sea surface temperature. Our findings show that the simulated long-term changes 19 of SWSA precipitation are dominated by externally forced anomalous expansion 20 of the Southern Hemisphere Hadley Cell (HC) and a persistent positive Southern 21 Annular Mode (SAM) trend since the late 1970s. Long-term changes in the HC ex-22

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tent and the SAM show strong co-linearity. They are attributable to stratospheric
ozone depletion in austral spring-summer and increased atmospheric greenhouse
gases all year round. Future ssp585 and ssp126 scenarios project a dominant role
of anthropogenic forcings on the HC expansion and the subsequent SWSA drying, exceeding the threshold of extreme drought due to internal variability as soon
as the 2040s, and suggest that these effects will persist until the end of the 21st
century.

Keywords Subtropical Andes drying trend · Hadley Cell expansion · Decadal
 variability · External forcing · CMIP6 · Detection and attribution · Future
 scenarios

33 1 Introduction

The southwestern South America (SWSA) region encompasses the Andean Cordillera 34 and adjacent territories from the Pacific coast to the continental arid lowlands in 35 Argentina and south of the dry Altiplano. This region is characterized by a marked 36 precipitation gradient from < 100 mm in the north (25 - 28° S) to well over 2000 37 mm in the south (40 - 45° S). Precipitation primarily occurs during austral winter 38 (June-August, JJA) associated with passing fronts embedded in the mid-latitude 39 Westerly flow and enhanced by the orographic effect of the Andes Mountains 40 (Montecinos and Aceituno, 2003; Garreaud et al., 2013; Viale et al., 2019). Lati-41 tudinal variations in the southeastern Pacific anticyclone and the subtropical jet 42 stream modulate the seasonality of these fronts, which commonly form at the 43 poleward limit of the Southern Hemisphere (SH) Hadley Cell (HC) (Montecinos 44 and Aceituno, 2003; Barrett and Hameed, 2017). Closely linked to large-scale at-45 mospheric circulation, SWSA precipitation supports the many glaciers and lakes 46 in the Andes and contributes to the flow of major streams and rivers along the 47 Cordillera. Indeed, most rivers originate in the upper Andes, where precipitation 48 is comparatively higher than in adjacent territories (Masiokas et al., 2019). With 49 some regional differences, SWSA has undergone a striking drying trend since the 50

1980s (e.g., Garreaud et al., 2013, 2017, 2020; Boisier et al., 2016, 2018), with 51 marked glacier retreats and lake-area reductions without precedent over the last 52 millennium (Garreaud et al., 2017; Pabón-Caicedo et al., 2020). A robust drying 53 trend of around -28 mm per decade during the austral winter rainy season has been 54 found in the southern part of SWSA (Boisier et al., 2018). In the northern drier 55 regions, the larger amplitude of the high-frequency variability in austral summer 56 (December-February, DJF) and fall (March-May, MAM) complicates the detection 57 of consistent trends in precipitation. However, a reduction in river streamflows sug-58 gests that a drying tendency is also taking place in spring (September-October, 59 SON) and summer in the north sector of SWSA (Boisier et al., 2018). Although 60 they display lower drought than observed, model simulations for the historical pe-61 riod of the 5th phase of the Coupled Model Intercomparison Project (CMIP5) can 62 capture such drying tendency in response to external anthropogenic forcings (Vera 63 and Díaz, 2015; Boisier et al., 2018). Due to the important socio-economic and eco-64 logical impacts of changes in water resources in SWSA (CR2, 2015; Norero and 65 Bonilla, 1999; Rosegrant et al., 2000; Meza et al., 2012), it is critical to investigate 66 further and quantify the potential impacts of climate change in this region. 67

Low-frequency rainfall changes in SWSA are also associated with internal cli-68 mate variability. Indeed, during austral winter and spring, SWSA annual moisture 69 conditions are tightly linked to the southeastern Pacific sea surface temperature 70 (SST) variability (Rutllant and Fuenzalida, 1991; Garreaud et al., 2009; Quintana 71 and Aceituno, 2012; Boisier et al., 2016) related to the El Niño Southern Oscil-72 lation (ENSO), the leading mode of interannual climate variability in the Pacific 73 Ocean (e.g., McPhaden et al., 2006). Long-term ENSO fluctuations imprint a pan-74 Pacific pattern of coherent SST decadal variability (Newman et al., 2016) through 75 atmospheric teleconnections (e.g., Alexander et al., 2002). This pattern is known 76 as the Interdecadal Pacific Oscillation (IPO) and is the leading mode of inter-77 nal, decadal to multidecadal variability in the Pacific Ocean (Folland et al., 1999; 78 Meehl and Hu, 2006). It is also typically referred to as its manifestation in the 79

wintertime North Pacific SST, the so-called Pacific Decadal Oscillation (Mantua 80 et al., 1997). During the positive (negative) phase, the IPO is characterized by 81 an ENSO-like pattern of warm (cold) SST anomalies across the tropical Pacific, 82 which extends in the subtropics over the eastern boundaries of the Pacific Ocean 83 (Trenberth and Hurrell, 1994; Meehl et al., 2009). The IPO has experienced a 84 trend from a positive (i.e., El Niño-like) to a negative (La Niña-like) phase over 85 the 1980-2014 period associated with an anomalous southward shift and spin-up of 86 the southeastern Pacific anticyclone (Jebri et al., 2020) and of the mid-latitudinal 87 storm-tracks over SWSA during the rainy season (Quintana and Aceituno, 2012; 88 Boisier et al., 2016, 2018). Thus, the concurrent IPO shift from positive to nega-89 tive phase might have contributed to the current prevailing SWSA dry conditions 90 (Masiokas et al., 2010; Quintana and Aceituno, 2012; Boisier et al., 2016). How-91 ever, during DJF and MAM seasons, rainfall interannual-to-decadal fluctuations 92 are also significantly influenced by the dominant mode of atmospheric circulation 93 variability in the mid-latitudes of the SH, the Southern Annual Mode (SAM), 94 also known as the Antarctic Oscillation (Gong and Wang, 1999; Thompson and 95 Wallace, 2000; Thompson et al., 2000). During SAM positive phases, a zonally 96 symmetric atmospheric pressure gradient is observed with negative and positive 97 anomalies over Antarctica and the mid-latitudes, respectively, favoring an anoma-98 lous poleward shift of the circumpolar westerlies and more constrained zonal flow 99 in mid-latitudes (Thompson and Wallace, 2000; Thompson et al., 2000), resulting 100 in less frontal rainfall in SWSA (Garreaud et al., 2009). These symmetric features 101 and the impact on SWSA precipitation are reversed during the negative SAM 102 phase. 103

The SWSA drying trends over recent decades could also result from external forcings. For instance, an intensification of the global hydrological cycle is projected under global warming due to a more extensive water vapour loading of the atmosphere, resulting in enhanced E-P (evaporation minus precipitation) in evaporative regions, and reduced E-P in precipitative regions according to the "wet-get-wetter,

dry-get-drier" paradigm (Helpd and Soden, 2006; Seager et al., 2010). Several ob-109 servational datasets, CMIP5 models and reanalyses also suggest an essential role 110 of externally forced large-scale dynamical changes. Observed southward shifts of 111 the subtropical drying regions and mid-latitudes baroclinic eddies during late aus-112 tral spring and summer have been associated with the SAM positive trend and 113 the expansion of the HC in recent decades (e.g., Gillett and Thompson, 2003; Pre-114 vidi and Liepert, 2007; Quintana and Aceituno, 2012; Boisier et al., 2018). The 115 HC expansion and SAM positive trends have both been attributed to increasing 116 atmospheric greenhouse gases (GHGs) concentration and stratospheric ozone de-117 pletion (e.g., Polvani et al., 2011; Kim et al., 2017; Jebri et al., 2020). However, the 118 tendency toward more La Niña events in relation to the recent IPO trend may also 119 contribute to the HC poleward shift (Nguyen et al., 2013; Allen and Kovilakam, 120 2017) and to more frequent positive SAM phases (Carvalho et al., 2005; Fogt et al., 121 2012). The counteracting effect of the recent ozone recovery (Eyring et al., 2010) 122 and the intensification of GHGs effect (Andreae et al., 2005) is a source of uncer-123 tainty for the understanding of the future state of these modes (Fogt and Marshall, 124 2020). 125

To our knowledge, previous studies have not explicitly examined (i) the re-126 spective contribution of anthropogenic forcing and internal climate variability on 127 the decadal-to-longer term variance of SWSA rainfall throughout the last century 128 and a half and (ii) the specific role of direct thermodynamical and dynamical 129 changes in SWSA related to the HC expansion and SAM trend. These are two 130 major points that are addressed in this study along with (iii) the attribution of 131 the modulation of SWSA precipitation to specific sources of external forcing over 132 the last decades (iv) and its projected changes for the 21st century. Here, we as-133 sess the SWSA hydroclimate changes over the last century using observations, 134 reanalyses and sets of 20 to 32 member-ensemble simulations conducted with the 135 stand-alone LMDz6A-LR atmospheric component (Hourdin et al., 2020) and the 136 IPSL-CM6A-LR coupled model (Boucher et al., 2020) as part of the 6th phase 137

of the CMIP exercise (CMIP6; Eyring et al., 2016). The paper is organized as
follows: in the next two sections, the data and methods used are introduced, the
results obtained are presented in section 4 and discussed in section 5. A summary
and the main conclusions are provided in section 6.

142 2 Data

¹⁴³ 2.1 The atmospheric and coupled model

We used the CMIP6 version of the Institut Pierre- Simon Laplace (IPSL) stand-144 alone atmosphere model, called LMDz6A-LR (Hourdin et al., 2020) and the cou-145 pled atmosphere-ocean general circulation model (GCM), called IPSL-CM6A-LR 146 (Boucher et al., 2020). LMDz6A-LR is coupled to the ORCHIDEE (d'Orgeval 147 et al., 2008) land surface component, version 2.0. In IPSL-CM6A-LR, the LMDz6A-148 LR is coupled to the oceanic component Nucleus for European Models of the 149 Ocean (NEMO), version 3.6, which includes other models to represent sea-ice 150 interactions (NEMO-LIM3; Rousset et al., 2015) and biogeochemistry processes 151 (NEMO-PISCES; Aumont et al., 2015). LR stands for low resolution, as the at-152 mospheric grid resolution is 1.25° in latitude, 2.5° in longitude and 79 vertical 153 levels (Hourdin et al., 2020). Compared to the 5A-LR model version and other 154 CMIP5-class models, IPSL-CM6A-LR was significantly improved in terms of the 155 climatology, e.g., by reducing overall SST biases and improving the latitudinal 156 position of subtropical jets in the SH (Boucher et al., 2020). The IPSL-CM6A-LR 157 is also more sensitive to CO₂ forcing increase (Boucher et al., 2020) and repre-158 sents a more robust global temperature response than the previous CMIP5 version 159 consistently with current state-of-the-art CMIP6 models (Zelinka et al., 2020). 160

¹⁶¹ 2.2 Experimental protocol

¹⁶² This study is based on a set of climate simulations generated as part of CMIP6

¹⁶³ (Table 1). We relied on the pre-industrial control (piControl) coupled run with

6

the external radiative forcing fixed to pre-industrial values for a measure of the
internal climate variability generated by the IPSL-CM6A-LR model.

We also used the 32-member ensemble of simulations for the historical period 166 (1850-2014), which is branched on random initial conditions from the piControl 167 168 run to ensure that ensemble members are largely uncorrelated during the simulation. All 32 members of this ensemble (referred to as the historical ensemble) 169 use the historical natural and anthropogenic forcings following CMIP6 protocol 170 (Eyring et al., 2016). They include concentrations of GHGs from 1850 to 2014 171 provided by (Meinshausen et al., 2017) while the standard CMIP6 tropospheric 172 and stratospheric ozone concentration were obtained from the Chemistry-Climate 173 Model Initiative (Checa-Garcia et al., 2018). Tropospheric aerosols, from natu-174 ral and anthropogenic sources (Hoesly et al., 2018; van Marle et al., 2017) are 175 also included along with historical stratospheric natural forcings, corresponding 176 to spectral solar irradiance-stratospheric ozone cycles (Matthes et al., 2017) and 177 the main volcanic eruptions prescribed with the aerosol optical depth (Thomason 178 et al., 2018). 179

Detection and attribution 10-member ensembles (Gillett et al., 2016) are also utilized to understand the role of the different external forcing components in the context of climate change. These experiments are equivalent to the historical ones but include each forcing individually namely either GHGs (hist-GHG), stratospheric ozone depletion (hist-stratO3), aerosols (hist-aer) and natural (hist-nat) while maintaining other forcing at their 1850 level.

Two 6-member ensembles for the 21^{st} century (2015-2100) future projections scenarios branched on randomly selected historical members at the year 2014 are also analyzed (O'Neill et al., 2016): the ssp585 equivalent to ~8.5 W m⁻² increased radiative forcing in the year 2100 due to GHG emissions (the highest future pathway across all CMIP6 scenarios) and the ssp126 mitigation scenario, which assumes the lowest of plausible radiative forcing effects (2.6 W m⁻²) in 2100 assuming substantial mitigation against global warming. Finally, we also use AMIP historical (amip-hist) simulations, which use the same external forcings since 1870 as the historical experiment but with imposed observed monthly SST on the LMDz6A-LR atmospheric component. Further details about external forcings implementation strategies in the experiments mentioned above are given in Lurton et al. (2020).

¹⁹⁸ 2.3 Observational data

In addition to model outputs, different observational products are used: in situ 199 and gridded observations data sets and reanalyses (Table 2). We analyze in situ 200 monthly precipitation data from 1960 to 2017 from rain gauges in 129 stations 201 located along the Andes, between 70° - 73° W and 20.5° - 46.5° S provided 202 by the Chilean Center for Climate and Resilience Research, Climate Explorer 203 (http://explorador.cr2.cl) (Fig. 1). Note that the station coverage is relatively 204 sparse outside the 27° - 43° S region in SWSA. We also rely on gridded products 205 of monthly precipitation provided by the Global Precipitation Climatology Centre 206 (version 2018; GPCCv2018) and the University of Delaware Air Temperature and 207 Precipitation (version 5.01; UDEL5.01), each of which using different interpola-208 tion methods. Note in Fig. 1 that the data coverage of the observations used for 209 the reconstruction of these gridded products is consistent with that of the *in situ* 210 observations, which are quite sparse outside SWSA in southern South America. 211

We use the Hadley Centre Sea Ice and Sea Surface Temperature version 1 212 (HadISST) gridded reconstruction of SST observations, which are used as bound-213 ary conditions in the LMDz6A-LR model amip-hist CMIP6 simulations. To study 214 the atmospheric dynamics, we also analyze the meridional wind and the sur-215 face pressure variables from several reanalyses: the NOAA-CIRES 20th Century 216 Reanalysis versions 2 (20CRv2), 2c (20CRv2c) and 3 (20CRv3), the ERA-20C, 217 ERA40 and ERA-Interim reanalyses of the European Centre for Medium-Range 218 Weather Forecasts and the National Center for Environmental Prediction (NCEP) 219

and the National Center for Atmospheric Research (NCAR) reanalysis version 1
(NCEP1) and 2 (NCEP2).

222 3 Methods

²²³ 3.1 Trend analyses and statistical testing

All anomalies are calculated by removing the monthly mean seasonal cycle of the 224 entire covered period (Table 1) from the time series before calculating seasonal av-225 erages (MAM, JJA, SON and DJF). To focus only on the variability from decadal, 226 interdecadal or longer time scales, time series are low-pass filtered (LPF) using a 227 Butterworth filter with a cut-off period of 8, 13 or 40 years, respectively (But-228 terworth, 1930). The linear trend for each month over 1979-2014 is obtained by 229 applying least squares on unfiltered yearly time series of monthly mean precipi-230 tation and represented in terms relative to the climatological mean (% decade⁻¹). 231 In this case, trend values are statistically tested with a Student t-test with the 232 number of degrees of freedom corresponding to the total number of years minus 233 one. To evaluate the ensemble-mean trend representativeness, we also indicate the 234 regions where at least 80% of the members display trends of the same sign than 235 the ensemble mean. The consistency between the anomalies of the climate param-236 eters and the modes of variability was estimated using the regression coefficient 237 (α) between these variables. When the time series are spatially distributed the re-238 sulting regression coefficients are presented as regression patterns. In this case, the 239 ensemble-mean regression coefficients are statistically tested using a random-phase 240 test, based on Ebisuzaki (1997), adapted to the regression (details in Villamayor 241 et al. (2018)). In turn, the level of uncertainty among individual members is indi-242 cated by showing the grid points, where at least 80% of them display a regression 243 coefficient of the same sign as the ensemble mean. When considering regionally 244 averaged time series for the ensemble mean, the uncertainty of the resulting re-245

²⁴⁶ gression coefficient is quantified by representing the range of values obtained with²⁴⁷ all members individually.

248 3.2 Climate indices

We consider five main sets of climate indices that are introduced below. Note that for ensemble simulations, the ensemble-mean index refers to the indices calculated from outputs previously averaged across all members. The 95% confidence interval among individual members, according to a Student t-test, is shown to quantify the uncertainty of the ensemble-mean indices .

254 3.2.1 SWSA precipitation indices

Empirical Orthogonal Function analysis of the 8-year LPF rainfall anomalies in southern South America $(25^{\circ} - 58^{\circ} \text{ S}, 65^{\circ} - 80^{\circ} \text{ W})$ is performed to isolate the leading mode of variability at decadal-to-longer time scales. The first Principal Component (PC1) for each season is standardized to get a measure of the variance. Regions with the highest variance are then used to build regional indices by averaging values over: $34^{\circ} - 49^{\circ} \text{ S}, 71^{\circ} - 77^{\circ} \text{ W}$ in MAM; $28^{\circ} - 40^{\circ} \text{ S}, 70^{\circ} - 76^{\circ}$ W in JJA; $30^{\circ} - 43^{\circ} \text{ S}, 70^{\circ} - 77^{\circ} \text{ W}$ in SON and $39^{\circ} - 50^{\circ} \text{ S}, 72^{\circ} - 78^{\circ} \text{ W}$ in DJF.

262 3.2.2 Global Warming and IPO sea surface temperature indices

A global warming index (GW) is calculated by spatially averaging 40-year LPF 263 annual-mean SST anomalies (SSTAs) over 45° S - 60° N. Residual SSTAs are then 264 derived by regressing out the GW SSTAs pattern. To analyze the IPO influence on 265 SWSA rainfall, an IPO index is computed using the residual SSTA following the 266 tripole approach of Henley et al. (2015), defined as the central equatorial Pacific 267 SST (10° S - 10° N, 170° E - 90° W) minus the average of northwest (25° - 45° 268 N, 140° E - 145° W) and southwest (15° - 50° S, 150° E - 160° W) Pacific SST. 269 The resulting IPO index is then LPF with a 13-year cut-off period. 270

271 3.2.3 Southern Annular Mode and Hadley Cell extent indices

Seasonal SAM indices are computed for each season as the standardized PC1 of the 8-year LPF sea level pressure (SLP) anomalies in the SH, south of 20° S. In order to describe the variability of the latitudinal position of the poleward edge of the HC in the SH, we define an index of the HC extent (HCE) for each season as the linearly interpolated latitude for which the 500-hPa meridional streamfunction (ψ_{500}) is equal to zero between 20° and 40° S.

²⁷⁸ 3.3 Decomposition of SWSA precipitation variance

A decomposition of the variance of SWSA precipitation into the components explained by the IPO and the GW SST indices is performed based on a multilinear regression analysis (Mohino et al., 2016). According to this, the total variance of SWSA precipitation (var[PR]) can be expressed in terms of the regression coefficients from the multilinear fitting that correspond to each index (α IPO and α GW, respectively) and a residual ($var[\varepsilon]$) as follows:

$$var[PR] = \alpha_{IPO}^{2} + \alpha_{GW}^{2} + var[\varepsilon]$$
⁽¹⁾

The residual term stands for the variance of the residual of the multilinear fitting, namely the variance that cannot be explained by the IPO and the GW indices. These three components are expressed as percentage of the total variance of SWSA precipitation.

289 3.4 Total moisture budget decomposition

A decomposition of the change in the net moisture budget, expressed as precipitation minus evaporation (P-E), into purely thermodynamic and dynamic components is performed in this work. According to the moisture budget equation, P-E equals the divergence of the moisture flux (Brubaker et al., 1993). The moisture flux is the mass-weighted and vertical integral of the product between the specific humidity (q) and the horizontal winds (**u**). Therefore, a change in P-E ($\delta(P-E)$) implies changes in q (i.e., thermodynamic changes (δTH) and **u** (i.e., dynamic changes (δDC). Considering such a change as an anomaly with respect to a reference period (denoted with subscript r), the thermodynamic and the dynamic components can be separated using the following approach (Seager et al., 2010):

$$\delta(P - E) = \delta T H + \delta D C + R E S \tag{2}$$

Following Ting et al. (2018), we define δTH and δDC as follows:

$$\delta TH \approx -\frac{1}{g\rho_w} \cdot \nabla \int_0^{p_s} \mathbf{u}_r \cdot \delta q \, dp \tag{3}$$

301

$$\delta DC \approx -\frac{1}{g\rho_w} \cdot \nabla \int_0^{p_s} \delta \mathbf{u}_r \cdot q \, dp \tag{4}$$

where g is the gravitational acceleration, ρ_w the density of water, p pressure levels and p_s the surface pressure.

The residual component (RES) mostly accounts for the moisture convergence changes due to transient eddies, but also includes nonlinear terms that are typically neglected in the approach of δTH and δDC (Seager et al., 2010; Ting et al., 2018).

307 3.5 Probability Density Functions of precipitation linear trend

To further evaluate the role of the IPO phase shift (such as during the 1979-2014 period) on SWSA precipitation trends in our model as compared to observations, we performed Probability Density Functions (PDFs) of the linear trend of SWSA precipitation for 36-year periods separately for ensemble members with a positive or negative IPO trend that is significant at the 5% level during the period (resulting in 25 members of each category). The two resulting PDFs are statistically compared by testing the null hypothesis that both present independent normal distributions with equal means and equal but unknown variances at the 20% significance level, according to a t-test. A t-test is also used to evaluate whether mean
trend values are significantly different from zero.

318 4 Results

319 4.1 Model validation

In this section, we present a comparison of the amip-hist and historical simulations with observational products to evaluate the ability of, respectively, the atmospheric component and the coupled model to simulate main aspects of precipitation in SWSA, such as climatology and tendency. We also check simulated IPO phase shifts and the emerging HC expansion, which are addressed in relation to the recent drying trend in SWSA.

326 4.1.1 SWSA precipitation

The SWSA region rainfall annual cycle over the period common to observations 327 and simulations (1960-2014) is represented in Fig. 2 (left panel). Considering the 328 spatial coverage of observations in SWSA (Fig. 1), the model validation is re-329 stricted to a latitudinal band between 20° - 47° S. This band is roughly 5° longi-330 tude width (i.e., two grid points in the model) and centered in 73.75° W. Most of 331 the rain gauge stations are located in the west Chilean territories adjacent to the 332 Andes where precipitation is enhanced by the orographic blocking effect on the 333 westerly atmospheric flow (Falvey and Garreaud, 2007; Smith and Evans, 2007; 334 Viale et al., 2019). The rain gauge observed annual cycle shows a latitudinal migra-335 tion of rainfall sustaining marked dry and wet seasons in austral summer (DJF) 336 and winter (JJA), respectively (Fig. 2a). The onset of the rainy season occurs 337 gradually during fall (MAM), with maximum rainfalls registered around 37° S in 338 June. This regional maximum is not well captured by the gridded observations, 339 most likely because of interpolation procedures applied to rain gauge data (Figs. 340

³⁴¹ 2c and 2e). The rainy season demise occurs in spring (SON), before minimum
³⁴² values are reached in summer.

The amip-hist and historical simulations reproduce the observed mean annual 343 cycle, with the onset in MAM, the rainy season in JJA with a maximum in June 344 around 37° S, the demise in SON and the dry season in DJF (Figs. 2g and 2i). 345 The model seems to overestimate climatological values north of 35° S and south 346 of 40° S. This discrepancy may instead be attributable to the lack of observations 347 at these latitudes. There is also a high level of uncertainties in observations re-348 garding the climatological amount. Annual rainfalls reach around 1460 mm in rain 349 gauge observations where it is most rainy between 35° S and 40° S, while gridded 350 products values are about two thirds of this amount (i.e., 910 mm in GPCCv2018 351 and 890 mm in UDEL5.01). Model ensemble-means for amip-hist and historical 352 simulations amount 1670 and 1450 mm respectively, with an ensemble standard 353 deviation of around 25 mm in both cases, which is close to in situ observations. 354 However, historical simulations generally give smaller values than amip-hist, espe-355 cially during the rainy season. This difference is independent of the size of both 356 ensembles (not shown) and most likely related to the influence of the observed 357 SSTs in amip-hist runs. 358

The annual cycle of the linear trend relative to the climatological mean over 359 the last 36 years (1979-2014) is represented in Fig. 2 (right panel). A significant 360 drying around 33° - 40° S in April-May and around 29° - 36° S in July is observed 361 consistently in all observations datasets (Figs. 2b,d,f). These drying trends indicate 362 a delay in the onset of the rainy season and less rainfall during the rainy season. 363 Although weak and not significant, negative trend values in September, together 364 with the features described before, suggest that, in general, there is a tendency 365 towards shorter and less effective rainy season in SWSA over the last decades. 366 The Hövmoller diagrams from gridded observations show positive trend during 367 late spring and summer south of 40° S, suggesting that the southernmost part of 368

³⁶⁹ SWSA has become wetter (Figs. 2d and 2f). However, this result is poorly reliable ³⁷⁰ due to the lack of observations in this part of the Andes (Fig. 1).

Both amip-hist and historical simulations indicate a widespread drying rela-371 tive to the simulated climatology along the year, with a maximum from around 372 $\sim 40^{\circ}$ S in DJFM to $\sim 27^{\circ}$ S in MJJAS (Figs. 2h and 2j), suggesting a shorten-373 ing and weakening of the rainy season. The drying trend is in general stronger 374 in amip-hist simulations than in historical runs (independently of the ensemble 375 size, not shown). All amip-hist members include the same observed SST, while 376 in historical simulation, SST variability is largely uncorrelated among ensemble 377 members. Averaging the historical ensemble allows therefore damping the influ-378 ence of SST on the simulated rainfall trends with respect to the role of external 379 forcing. Differences with amip-hist on the other hand, emphasize the contribution 380 of observed SST variability. The drying trend during the onset and demise of the 381 rainy season is more intense and significant in the amip-hist ensemble mean than 382 in the historical one. This suggests a dominant role of observed SST variability 383 (such as the shift toward a negative IPO) in the rainy season shortening during 384 1979-2014, probably amplified by external forcing influence. 385

To summarize, despite the constraints of comparing local observations with 386 gridded data from GCMs due to the complex orography of the region, the model 387 can reproduce the main observed rainfall climatological features and recent trends. 388 Differences between the amip-hist and the historical coupled simulations suggest 389 a combination of internal and external factors in driving the SWSA drying trend. 390 Previous generation CMIP5 models show an overall underestimation of the SWSA 391 precipitation response to external forcing (Vera and Díaz, 2015; Boisier et al., 392 2018) that is coherent with the IPSL-CM6A-LR simulations. On the other hand, 393 a preliminary CMIP6 multi-model study places the IPSL-CM6A-LR among the 394 GCMs that best represent the atmospheric changes over southeast Pacific that 395 modulate SWSA rainfall during the last century (Rivera and Arnould, 2020), which 396

³⁹⁷ supports the use of this model to study the long-term precipitation variability in ³⁹⁸ this region.

399 4.1.2 IPO phase shifts

We also evaluate the model ability to simulate IPO phase shifts comparable to the 400 observed shift over the 1979-2014 period addressed by Boisier et al. (2016) in terms 401 of the amplitude of the associated SST anomalies. To this aim, Fig. 3 represents 402 IPO indices from observational data (HadISST1) over 1870-2014 and from the 403 piControl run over a representative period of equal length, as well as the linear 404 trend values obtained along both IPO indices in centered running windows of 15-40 405 years long. The trend graphics show blue and red colored plumes, corresponding 406 to negative and positive trend values, respectively. The plumes resulting from the 407 piControl IPO index are, in general, narrower than those from observations. This 408 reveals that the model underestimates the observed persistence of the IPO phases. 409 However, the model succeeds in simulating 36-year IPO phase shifts of up to -0.2 410 °C per decade comparable to the one observed over 1979-2014. 411

412 4.1.3 Hadley Cell expansion

To evaluate the model's ability to reproduce the HC expansion in recent decades, 413 we compare the linear trend of the seasonal HCE indices calculated with eight 414 different reanalyses and the amip-hist and historical simulations for their common 415 23-year period (1979-2001) and for a 40-year period (1971-2010), common to the 416 simulations and five reanalyses (Fig. 4). The trend values obtained over the short 417 period show large dispersion between the different members of the amip-hist and 418 historical simulations, compared with those of the longer period. This shows the 419 dependence of the tendency of the HCE index on short-term stochastic internal 420 variability. In the longer term, the trend is more robust across model members 421 suggesting an influence of external forcings. Comparing across seasons, the simu-422 lated HCE trend is better constrained among members in JJA and most uncertain 423

in DJF, when the HC is most variant and presents wider expansion. Regarding
observations, the trend values obtained from reanalyses show large spread as well.
Despite the large uncertainty, when the 40-year long period is considered, there is
solid agreement among reanalyses regarding the Southern HC expansion towards
the pole in recent decades (Grise et al., 2019).

The model simulates trends that are within the range of those in the reanal-429 yses. The trend of the simulated HCE in the ensemble mean is negative in all 430 seasons and in both amip-hist and historical simulations, being widest in DJF. In 431 MAM, the poleward expansion of the HCE is the weakest and close to zero when 432 simulated by the historical simulation, while the amip-hist simulation shows a 433 broader expansion. In contrast, during the other seasons, the simulated ensemble-434 mean expansion is similar in both simulations, being wider in the historical over 435 the 40-year-period. These results suggest that the external forcing mostly induces 436 the HC poleward expansion in JJA, SON and particularly in DJF, but in MAM 437 this effect is weak and, presumably, the influence of SST internal variability is also 438 relevant. 439

440 4.2 Relative roles of forced versus internal variability on SWSA precipitation
441 decadal variability and trend

In this subsection, we focus on the amip-hist ensemble to unravel the contribution 442 of observed SST to SWSA precipitation low-frequency variability. The first PCs 443 of the 8-year LPF seasonal precipitation anomalies account for much of the total 444 decadal-scale rainfall variability in southern South America ($25^{\circ} - 58^{\circ}$ S; $65^{\circ} - 80^{\circ}$ 445 W). The explained total variance of the PC1 is 59.1% in MAM, 75.7% in JJA, 446 67.0% in SON and 69.9% in DJF. These indices together with their respective 447 regression patterns show that most of the precipitation variability in southern 448 South America is concentrated in a dipole of opposite anomalies between the 449 middle $(25^{\circ} - 45^{\circ} \text{ S})$ and high $(> 50^{\circ} \text{ S})$ latitudes in SWSA (Fig. 5). The seasonal 450 patterns in DJF and MAM also present significant anomalies in subtropical South 451

⁴⁵² America east of the Andes showing an opposite sign to those recorded in middle
⁴⁵³ SWSA latitudes.

These regression patterns allow identifying where the most substantial low-454 frequency variability occurs in the amip-hist simulations (boxed areas in Fig. 5). 455 These regions vary depending on the season, showing a north-to-south shift from 456 winter (JJA) to summer (DJF), respectively, and are co-located with the centers 457 of maximum rainfall linear trends (gray contours in Fig. 5). These regions with 458 the most substantial negative rainfall anomalies are associated with PC1s positive 459 long-term trend (Fig. 5, left panels), which accounts for nearly the total tendency 460 of the area-averaged precipitation anomalies for each season (Table 3). From 1970 461 to 2014 a deficit of 6.2 mm per decade and around 15.2 mm per decade is simulated 462 for the rainy (JJA) and dry summer (DJF) seasons respectively. 463

Previous studies have shown evidence for the influence of SST internal vari-464 ability and external forcings on SWSA rainfall recent changes (e.g., Boisier et al., 465 2016, 2018). In line with these previous works and to identify potential connections 466 between simulated low-frequency variations of SWSA precipitation and observed 467 internal modes of climate variability, we regress PC1s on anomalies of SST, SLP 468 and wind (Fig. 6) before and after removing the long-term trend signal in PC1s 469 with a 3rd-degree polynomial fit (dotted lines in Fig. 5). Such detrending will allow 470 emphasizing typical SST and teleconnection patterns related to internal variability 471 modes. 472

Without detrending, SST regression patterns show warm anomalies almost 473 globally distributed except in the tropical Pacific where an IPO-like relative cool-474 ing dominates all year round (left panel in Fig. 6). The warm pattern is most 475 widespread in DJF, with strong anomalies over western Pacific, the Indian and 476 Atlantic Oceans, especially in the southern basin, and weakest in JJA. In turn, 477 the tropical Pacific cooling is more intense in SON and JJA and almost negligible 478 in DJF and MAM. This global SST pattern is reminiscent of an IPO negative 479 phase combined with the global warming signal induced by external forcing. In 480

⁴⁸¹ DJF and MAM surface winds and SLP anomalies show a poleward shift of the ⁴⁸² westerlies and a SLP maximum slightly south of 30° S associated with the SWSA ⁴⁸³ drying. This suggests a link with the strengthening of the SAM with a weaker ⁴⁸⁴ influence of the negative IPO pattern on the SWSA drying in austral summer and ⁴⁸⁵ fall, in agreement with Boisier et al. (2018).

Regression patterns obtained with the detrended PC1s highlight a sea-level 486 pressure high and significant anomalous easterlies in Southeastern Pacific in JJA 487 and SON, suggesting an atmospheric teleconnection with a negative IPO-like SSTA 488 pattern most prominent in austral winter and spring (right panel in Fig. 6). Ex-489 tratropical SLP anomalies are less zonally symmetric than with the non-detrended 490 PC1s with an anomalous jet of westerlies passing through SWSA south of around 491 45° S, which is coherent with the anomalous SLP gradient. Therefore, this pattern 492 may suggest that, in response to a negative IPO, there is a poleward shift of the 493 storm-tracks embedded in the zonal flow (Garreaud et al., 2013), resulting in less 494 intrusion of humid air masses over middle latitudes in SWSA. Comparison of re-495 gression patterns before and after PC1s detrending indicates therefore that precip-496 itation trends in winter (JJA) and spring (SON) are largely influenced by internal 497 IPO related SST variability (Boisier et al., 2016). On the other hand, high-latitude 498 atmospheric circulation associated with the SST warming signal dominates SWSA 499 rainfall trends especially in summer (DJF) and fall (MAM) (e.g., Boisier et al., 500 2018). 501

To further evaluate the respective roles of the global warming and internal SST 502 variability on SWSA precipitation, a multilinear regression analysis over 1870-503 2014 is performed using as predictands the simulated seasonal indices of SWSA 504 precipitation (i.e., area-weighted mean over boxed areas in Fig. 5) and the IPO 505 and GW observed SST indices as predictors (Fig. 7). The results show that the 506 IPO barely explains 19% and 15% of the precipitation low-frequency total variance 507 in JJA and SON, respectively, and that its contribution is almost null in MAM 508 and DJF (Fig. 7). The GW index explains a larger proportion of rainfall variance 509

in all the seasons except in JJA. The external forcing influence is exceptionally 510 remarkable in DJF, where the GW dominates the precipitation variance above the 511 IPO, which shows very little impact. However, the explained variance by the GW 512 index only represents the indirect external forcing influence on SWSA rainfalls 513 through induced SST anomalies. It does not account for the externally forced 514 changes in atmospheric dynamics and direct thermodynamic effects in SWSA. 515 The fact that the residual component is considerably large in all seasons suggests 516 that SST changes are not a significant driver of the SWSA precipitation variability 517 at decadal-to-longer time scales. 518

A possible explanation for the low response of the simulated SWSA precipita-519 tion to the IPO could be due to the unrealistically weak atmospheric teleconnection 520 simulated by the model. However, the SLP anomalies regressively associated with 521 the IPO in JJA over southeastern Pacific (blue box in Fig. 6b) in HadSLP2 ob-522 servations (0.24 hPa per standard deviation) lies within one standard deviation 523 of the ensemble-mean values of the amip-hist and historical simulations (respec-524 tively, 0.32 ± 0.14 and 0.34 ± 0.15 hPa per standard deviation). Hence, we cannot 525 attribute the relatively weak influence of the internal SST decadal variability on 526 simulated rainfall to insufficient sensitivity of the atmospheric component of the 527 model to SST anomalies. 528

The SWSA precipitation response to the IPO simulated in amip-hist, histori-529 cal and piControl simulations (see Table 1) are shown in the bottom panel of Fig. 530 7. MAM and DJF rainfall responses are positive in the amip-hist and historical 531 ensembles, but weak and not emerging from the internal variability as represented 532 by the piControl simulation. In turn, this positive signal is significant in amip-hist 533 in JJA and SON, the strongest being in JJA. In the historical simulations, the 534 response to the IPO is also positive and strongest in JJA but not significant as 535 in all seasons, consistently with the unforced piControl run. These observation 536 suggest that the difference in the SWSA response to IPO between forced and un-537 forced coupled model simulations is not significant. Therefore, according to these 538

results, it can be inferred that the IPO can impact on the simulated SWSA precipitation. Still, its influence is weak as compared to external forcing and only robustly reproducible with the SST-forced amip-hist simulations in JJA and, to a lesser extent, in SON. Next, we examine the relative influences of dynamical and thermodynamical external forcings on the low-frequency variability and trend of SWSA rainfalls.

545 4.3 Role of dynamical vs. thermodynamical changes in SWSA

In the previous subsection, anomalous poleward shift of the zonal circulation in 546 high-latitudes in the SH is attributed to external forcings. This suggests that ex-547 ternal forcings act on SWSA precipitation through dynamical changes. However, it 548 has been shown that long-term dynamical changes cannot account for all the exter-549 nally forced trend in subtropical precipitation with direct thermodynamic effects 550 playing an important role (Schmidt and Grise, 2017). Therefore, the question arises 551 as to whether the forced component of precipitation in SWSA is mostly induced 552 by dynamic or thermodynamic processes. To shed light on the leading process, we 553 express the change in the net surface moisture budget as the difference of precipita-554 tion minus evaporation (P-E) over 2005-2014 versus 1851-1910 ($\delta(P-E)$). Besides, 555 we decompose P-E into a thermodynamic component (δTH), due to changes in the 556 specific humidity, a dynamic component (δDC) due to changes in circulation, and a 557 third component associated with transient eddies (Seager et al., 2010). The change 558 of the annual mean P-E obtained from the ensemble-mean historical simulations 559 roughly presents a hemispheric pattern that represents the "wet gets wetter and 560 dry gets drier" paradigm of Helpd and Soden (2006) (Fig. 8a). It is worth noting 561 that the deficit of moisture supply in SWSA stands out across other extratropical 562 continental regions of the SH. The decomposition of the change of seasonal mois-563 ture supply (Fig. 8b), reveals that changes in the atmospheric circulation account 564 for a large share of the total change compared to the direct thermodynamic effects 565 of external forcing in all seasons. Consistent with these findings, in the next sec-566

tion, we examine the potential large-scale dynamics sources associated with the
 induced SWSA drying.

⁵⁶⁹ 4.4 Connections between the HCE, SAM and SWSA rainfall

An anomalous shift of the circumpolar circulation in the SH high-latitudes, like 570 the one shown by the regression patterns of the undetrended PC1s of SWSA pre-571 cipitation on circulation forcings (Fig. 6), has been related to a persistent positive 572 SAM trend and a widening of the HC in response to external forcings (e.g., Gillett 573 and Thompson, 2003; Amaya et al., 2018). The historical ensemble mean shows 574 that the seasonal low-frequency indices of SWSA precipitation, HCE and SAM 575 describe a consistent trend (left panel in Fig. 9; note that the SAM indices are 576 reversed). This trend is stronger since $\sim 1970s$ and more pronounced in DJF and 577 MAM than in JJA and SON. The square of the correlation coefficient (R^2) reveals 578 high co-linearity among the three indices in all seasons (right panel in Fig. 9). 579 SWSA precipitation is almost equally correlated to HCE and SAM, showing the 580 highest co-linearity in DJF ($R^2 = 0.87$ and $R^2 = 0.92$, respectively) and the lowest 581 in JJA ($R^2 = 0.73$ and $R^2 = 0.74$, respectively). In turn, the HCE and the SAM 582 also show a strong co-linearity between each other that is strongest in DJF ($R^2 =$ 583 0.95) and weakest in JJA ($R^2 = 0.86$). This result suggests that both dynamical 584 large-scale atmospheric modes vary in synchronously and modulate SWSA precip-585 itation in the same direction in response to external forcings. Since both HCE and 586 SAM are highly co-linear, we next focus only on HCE index to investigate and 587 attribute the simulated trends. 588

589 4.5 Attribution of forced variability

⁵⁹⁰ Until the early 1970s, the HCE indices of all coupled simulation oscillate around ⁵⁹¹ the piControl climatological value and within the threshold of internal variations in ⁵⁹² all seasons (Fig. 10, left panel). Afterwards, the historical ensemble-mean indices show a poleward expansion in all seasons, even exceeding the bounds of internal variations by the 2000s in JJA, SON and DJF. At the same time, the hist-nat experiment simulates no appreciable change of the HCE variability. Such tendencies evidence the role of the anthropogenic forcings on the recent HC expansion.

To highlight the role of external forcings, the linear trend over 1970-2014 of 597 the ensemble-mean HCE indices is analyzed (Fig. 10, middle panel). The trend 598 displayed by the historical simulation is negative, denoting a poleward expansion, 599 and significantly different from zero (with 99% confidence interval) in all seasons, 600 being even robust among all members in DJF. The expansion is widest in this 601 season with a trend of -0.20 °lat per decade, then -0.04 °lat per decade in MAM, 602 -0.07 °lat per decade in JJA and -0.11 °lat per decade in SON. Large error bars 603 evidence the great influence of internal weather noise. GHG forcing alone has a 604 large effect all year round, with a signature that emerges out of the internal noise 605 by the early 2010s in MAM and the 2000s in JJA and SON, with significant 606 1970-2014 linear trends of -0.06, -0.09 and -0.10 $^{\circ}$ lat per decade, in the respective 607 seasons. Such trend values highlight the leading role of GHGs in the HC expansion 608 during these seasons. In DJF, instead, the stratospheric ozone depletion leads the 609 HC expansion, inducing wider shift (-0.11 °lat per decade) than GHGs (-0.06 °lat 610 per decade) separately. Anthropogenic aerosols (aer) induce a significant HCE 611 equatorward shift in JJA (0.02 °lat per decade) and SON (0.06 °lat per decade) 612 from around the mid-1990s with a peak in the mid-2000 and a partial recovery 613 afterwards. Nevertheless, this contraction is constrained within the threshold of 614 internal variability. The 1970-2014 linear trend of the "sum" index, obtained by 615 adding the ensemble-mean HCE anomalies from individual forcings, is significantly 616 close to the historical one (with 95% confidence interval) in DJF, MAM and JJA. 617 Therefore, the effects of the external forcings are, overall, additive all year round, 618 except in SON due to the influence of anthropogenic aerosols that tends to offset 619 the trend of the "sum" index (-0.04 $^{\circ}$ lat per decade) respectively to historical 620 experiments (-0.11 °lat per decade). 621

To quantify the role of the HCE in the forced SWSA drying over 1970-2014, 622 we represent the total trend of SWSA precipitation from the ensemble-mean sim-623 ulations and the HCE-coherent trend in Fig. 10 (right panel). The historical ex-624 periment shows a precipitation reduction of 4.5, 4.4, 3.6 and 10.3 mm per decade 625 in MAM, JJA, SON and DJF, respectively, in response to all external forcings. 626 The HCE-coherent trend underestimates by 40% the total drying in MAM. Still, 627 it shows no significant difference with the total trend in the other seasons, which 628 highlights the overall leading role of the HC expansion on the drying trend induced 629 by external forcing. In MAM, in response to only GHGs the total precipitation 630 trend is significantly similar to the one induced by all forcings and is strongly asso-631 ciated with the HC expansion. The rest of individual forcings generate no signifi-632 cant trends. In JJA and SON, GHGs alone induce a drying trend that respectively 633 doubles and equals that caused by all forcing together and that is tightly linked 634 to HC expansion. In both seasons, ozone depletion also contributes to drying, but 635 its impact is not associated with the HCE shift. In turn, aerosols and, to a lesser 636 extent, natural forcings partially counteract the drying. Aerosols positive effect 637 on precipitation is consistent with simultaneous HC contraction in these seasons. 638 Nevertheless, the HCE index barely accounts for 23% and 18% of the total trend 639 induced by aerosols respectively in JJA and SON. In DJF, the effect of GHGs 640 and ozone depletion separately induce each around half of the trend displayed by 641 the historical simulation, while the other forcings induce no significant trend. The 642 HCE-coherent trend underestimates by 60% the drying attributed to GHGs but 643 accounts for the entire trend induced by ozone depletion. The difference between 644 the ensemble-mean trend of the "sum" index of SWSA precipitation and the one 645 from the historical simulation suggest that the effects of individual forcings are 646 barely additive. Nevertheless, this difference is not statistically significant with 647 high significance interval, due to the considerable uncertainty among individual 648 members. 649

650 5 Discussion

⁶⁵¹ 5.1 The role of the IPO

Our results highlight the leading role of the SAM and the HCE on the recent 652 SWSA long-term drying trend in response to anthropogenic forcings. This is in 653 line with Boisier et al. (2018)'s finding showing that GHGs and the ozone de-654 pletion drive the SWSA drying trend over 1960-2016 in association with positive 655 SAM anomalies. Nonetheless, Boisier et al. (2016) concludes that the forced dry-656 ing trend at multidecadal scale can be substantially modulated by internal SST 657 variability through the IPO. This previous work focuses on the 1979-2014 period 658 and concludes, based on a linear regression model, that approximately 40% of the 659 SWSA drying trend is attributable to the concurrent pronounced negative trend 660 in the IPO index (Fig. 3). This statement opposes our result on the decomposition 661 of the decadal variance of SWSA precipitation, which reveals little influence of the 662 IPO compared to external forcing (Fig. 7). 663

To shed light on this discrepancy, we focus on the IPO-coherent trend of pre-664 cipitation over 1979-2014 using the ensemble-mean amip-hist precipitation (upper 665 panel in Fig. 11). This trend exceeds the one attributable only to the internal vari-666 ability of observed SST (i.e., the trend of ensemble-mean amip-hist precipitation 667 minus the trend from the historical ensemble mean) in all seasons. This diagnostic 668 evinces that a linear regression model can mislead the IPO signal on the SWSA 669 precipitation with the trend associated with external forcing. On the other hand, 670 the amip-hist ensemble-mean 1979-2014 drying trend notably exceeds the forced 671 component represented by the ensemble mean of the historical coupled runs. This 672 may suggest that the linear evolution of observed SST in 1979-2014, character-673 ized by a steep negative IPO phase-shift, has a major contribution to the drought 674 along with external forcing. Nevertheless, the amip-hist ensemble-mean precipita-675 tion trend may also represent an amplification of the forced drying by non-linear 676

air-sea interactions and atmospheric processes rather than the signature of the specific IPO phase-shift.

To further assess whether the IPO phase shifts can directly induce decadal 679 trends in the simulated SWSA precipitation, we perform PDFs of the precipita-680 tion trend over 1979-2014 and across all other possible 36-year periods since 1870 681 (bottom panel in Fig. 11). We use all the amip-hist members separately and rep-682 resent the precipitation trend only in 36-year periods in which the IPO presents a 683 phase shift and significant linear trend. 50 IPO trend values are obtained, equally 684 distributed between positive and negative shifts ranging between $\pm 0.09^{\circ}$ C per 685 decade centered on respective means of 0.17° C and -0.16° C per decade. Then 686 the precipitation trend values are classified according to whether the shift is from 687 a negative to a positive IPO phase (red bars) or vice versa (blue bars). Accord-688 ing to a two-sample t-test, the null hypothesis that the red and the blue PDFs 689 show normal distributions with equal means and variances cannot be rejected, 690 considering low confidence intervals (p>0.92) in all seasons. Their means, repre-691 sented by the red and blue solid vertical lines, show weak changes compared to 692 the ensemble-mean trend over 1979-2014 (green solid lines; corresponding to an 693 IPO shift of -0.21° C per decade). But they are negative in all cases, except in 694 MAM where the blue line indicates an almost null positive trend (0.2 mm per 695 decade). Therefore, these results suggest no significant relationship between the 696 linear trend of the simulated precipitation and the sign of the IPO phase shifts 697 over 36-year periods. 698

This conclusion is supported by similar PDFs performed using the 32 historical members (not shown). The resulting PDFs are not significantly different between each other and to the previous ones from the amip-hist simulations, according to a two-sample t-test. In addition, the resulting mean trends (red and blue dashed vertical lines in Fig. 11) are weaker than the drying represented by the historical ensemble mean over 1979-2014 in all seasons (green dashed lines) but negative ⁷⁰⁵ in most cases, consistently with the amip-hist simulations, or positive but almost⁷⁰⁶ zero in JJA.

In contrast, the same analysis applied to trends over 20-year or shorter peri-707 ods across the amip-hist simulations reveals a break up between the red and blue 708 PDFs (not shown). They show prevailing positive (negative) rainfall trends con-709 current with positive (negative) IPO phase shifts. This means that the amip-hist 710 simulations can reproduce SWSA precipitation trends in response to IPO phase 711 shifts at intradecadal-to-decadal time scales, in agreement with other works based 712 on instrumental data (Masiokas et al., 2010). However, at multidecadal-to-longer 713 timescales, such as the 36-year-long negative IPO trend observed since the 1979, 714 the simulated effect of the IPO over the last century and a half vanishes and 715 becomes insignificant compared to the leading role of external forcings. 716

In summary, our results show that the IPO explains little variance of SWSA precipitation multidecadal variability over the last century and a half, compared to external forcings. However, during 1979-2014 the simulated SWSA drying trend attributable only to internal variability of observed SST accounts for a large part of the total trend (Fig. 11). Therefore, we can conclude that the IPO is not a leading cause of long-term variability of SWSA precipitation, but it can contribute to amplify forced drying trends.

It has to be considered that the methodology used in this work is conditioned by the IPSL-CM6A-LR model biases in the atmospheric circulation. Therefore, the IPO influence on SWSA precipitation long-term variability should be assessed in the multi-model framework of CMIP6 to strengthen this conclusion.

$_{728}$ 5.2 The simulated HC expansion

Regarding the simulated seasonality in HCE, model trend values are comparable
with those from observations, with simulations reproducing a poleward expansion
over recent decades. In agreement with Hu et al. (2011, 2018) and model-based
studies (Staten et al., 2012; Hu et al., 2013; Grise et al., 2018), the HCE expan-

sion shows strong seasonality, being larger in austral summer (DJF) than in other 733 seasons. Nevertheless, it is difficult to constrain a precise widening rate of the 734 HCE expansion due to the large spread among the trends from reanalyses and 735 the ensemble simulations. Regarding the wide range of trend values resulting from 736 the simulated ensemble members, it can be inferred that some part of the large 737 uncertainty is attributable to internal variability. In turn, observed and simulated 738 trends are comparable when the internal variability is considered, suggesting that a 739 large part of the observed HCE expansion is accounted for by internal atmospheric 740 variability (Grise et al., 2018). On the other hand, there is also large discrepancy 741 in the HCE expansion rates derived from the different reanalyses. Apart from dif-742 ferences associated with distinct assimilation methods and model biases, this large 743 discrepancy has been attributed, in part, to shortcomings regarding the conserva-744 tion of mass in the meridional mean circulation (Davis and Davis, 2018), especially 745 in old generation reanalyses (Grise et al., 2019). 746

747 5.3 HCE and SAM co-linearity

Our results show strong co-linearity between the SAM, HCE and SWSA precip-748 itation, evidencing that both modes induce changes on precipitation in the same 749 direction. Furthermore, the simulated SAM strengthening and the HCE expansion 750 are attributed to the same external forcings. This suggests a connection between 751 circulation changes in tropical and high latitudes through a link of both vari-752 ability modes under the effect of external forcing, in agreement with previous 753 works (Thompson and Wallace, 2000; Previdi and Liepert, 2007). GCMs simu-754 late an anomalous rise of the extratropical troposphere in association with the 755 SAM strengthening (Previdi and Liepert, 2007) and the expansion of the HCE 756 (Lu et al., 2007) in future projections that consider strong external forcing effects. 757 These anomalies are related to the increase of atmospheric static stability, which 758 in turn is associated with a poleward expansion of the tropospheric baroclinicity. 759 Subsequently, there is a shift in the same direction of the hemispheric circulation 760

⁷⁶¹ such as the westerly jets, which involves changes in both the SAM (Thompson ⁷⁶² et al., 2000) and the HCE (Staten et al., 2019), and in the mid-latitude storm-⁷⁶³ tracks modulating the SWSA precipitation. Nevertheless, the factors that control ⁷⁶⁴ the extratropical atmospheric static instability are still unknown and, by exten-⁷⁶⁵ sion, the mechanisms that explain the link between changes in the SAM and the ⁷⁶⁶ HCE.

⁷⁶⁷ 5.4 The role of anthropogenic forcing

Our results show that anthropogenic forcings have largely induced the recent 768 SWSA drying through dynamical changes associated with the HCE and SAM. 769 GHGs play the leading role on the HC expansion and the subsequent SWSA dry-770 ing trend all year round except in DJF. In this season, the effect of GHGs is 771 combined with the ozone depletion, which is the dominant factor. Although not 772 shown in this paper, the leading effects of GHGs and ozone depletion are found to 773 be similar on the HCE and the SAM. We also find that the effect of anthropogenic 774 aerosols can offset the simulated HC expansion and the SWSA drying trend over 775 1970-2014. This is most likely related to the counteracting effect of aerosols on the 776 global warming induced by GHGs (Andreae et al., 2005). However, the HC con-777 traction barely explains a fraction of the forced rainfall increase, which is largest in 778 JJA (Fig. 10) coincident with positive thermodynamics of the P-E imbalance (Fig. 779 8). This suggests that aerosols may induce precipitation changes in SWSA through 780 thermodynamical processes. Nevertheless, aerosols effects are highly uncertain and 781 poorly constrained by climate models (Andreae et al., 2005; Boucher et al., 2013; 782 Carslaw et al., 2013; Oudar et al., 2018), suggesting that further investigation 783 should be done to shed light on this concern. 784

⁷⁸⁵ 5.5 Projected of HCE and SWSA precipitation

Aerosol emissions are likely to decrease in the future intensifying the GHGs effect 786 (Andreae et al., 2005) concurrently with the stratospheric ozone recovery (Eyring 787 et al., 2010). Consequently, an acceleration of HC expansion associated with aerosol 788 depletion during JJA and SON could be expected in the coming years. At the same 789 time, a slowdown of HC would occur in DJF with ozone recovery. However, no solid 790 assumptions on the HCE future changes and the subsequent impacts on SWSA 791 precipitation can be made without taking into account the evolution of GHG 792 emissions. To address this concern, we analyze the ssp585 and the ssp126 future 793 projections, which incorporate the aforementioned future evolution of the aerosols 794 and stratospheric ozone under business-as-usual and mitigation scenarios of GHG 795 emissions simultaneously (Lurton et al., 2020). The resulting ensemble-mean in-796 dices of HCE and SWSA precipitation are represented in Fig. 12, following those 797 of the historical simulations. The ssp585 projection shows a sustained poleward 798 expansion of the HCE in response to the substantial radiative forcing increase 799 until the end of the current century. This result is consistent with previous work 800 based on different GCMs (Grise et al., 2018; Staten et al., 2018). A SWSA drying 801 follows the HC expansion in all seasons. The projected drying exceeds the thresh-802 old of extreme drought due to internal variability by the 2040s in MAM and JJA, 803 the 2030s in SON. In DJF, the drying is shown to be extreme since the 2000s, 804 according to the historical simulation. The yearly precipitation over 2091-2100 is 805 projected to be 37% lower with respect to 1851-1910 along with an annual-mean 806 HCE poleward shift of 2.4 °lat. In turn, ssp126 experiment projects a 13% loss of 807 yearly precipitation and an annual-mean HC poleward expansion of 0.7 °lat. This 808 mitigation scenario represents a stabilization of the HC expansion and the SWSA 809 drying close to the threshold of extreme rates in all seasons rather than a recovery. 810 Under the ssp585 scenario, an intense moisture deficit is also projected by the 811 end of the 21st century in other subtropical regions, apart from SWSA (Fig. 13). 812 Coastal areas in southern Angola and Namibia, south of South Africa, western and 813

southeastern Australia also show negative P-E change associated with the HCE 814 evolution. The historical simulation, in contrast, does not show such strong link 815 in these regions (Fig. 8a). This suggests that the HC expansion could become an 816 essential modulator of precipitation in these mostly dry regions as in SWSA in a 817 future with high GHGs emissions. To further understand and attribute the future 818 evolution of precipitation in these regions, the contribution of the HC expansion 819 against thermodynamical direct effects and the relative role of external forcings 820 with respect to internal long-term variability should be addressed more thoroughly. 821

822 6 Conclusions

The IPSL-CM6A-LR model simulates an emerging long-term drying trend in 823 SWSA roughly since the early 1980s in response to the observed SST and external 824 forcings, consistent with observations and other GCMs (e.g., Vera and Díaz, 2015; 825 Boisier et al., 2018). A modulating effect of the internal SST variability on the 826 simulated SWSA drying over 1979-2014 is detected, related to the IPO as sug-827 gested by previous work (Boisier et al., 2016). However, the simulated impact of 828 the IPO on the overall decadal variability of SWSA precipitation across the last 829 century and a half is found to be weak and secondary compared to the effect of 830 external forcings. 831

The external forcing modulates the simulated SWSA precipitation indirectly 832 through dynamic changes, prevailing over direct thermodynamic effects. Specifi-833 cally, this work relates the simulated drought in SWSA with a concurrent strength-834 ening of the SAM and expansion of the HC, which occur in response to external 835 forcings. Simulated forced changes in these two variability modes act on SWSA 836 precipitation in the same direction and respond to the same components of ex-837 ternal forcings. Both modes react more strongly to external forcing in DJF than 838 in other seasons principally in response to the stratospheric ozone depletion and, 830 in second place, to the GHGs, in agreement with previous studies (e.g., Polvani 840 et al., 2011; Kim et al., 2017; Jebri et al., 2020). In the rest of the seasons, the 841

- ⁸⁴² ozone depletion effect is weaker and the GHG forcing prevails. Future projections
- ⁸⁴³ suggest that the GHG effect determines the HCE, and consequently, the long-term
- $_{\rm 844}$ $\,$ variability in SWSA precipitation throughout the 21 $^{\rm st}$ century.

Table 1 List of simulations with the IPSL-CM6A-LR model analyzed in the paper, the period covered, the size of the ensemble and imposed boundary conditions: observed SST, external forcing effects of GHGs, stratospheric ozone depletion, anthropogenic aerosols and natural forcings (namely stratospheric eruptions and spectral solar irradiance). (*) The forcing values in the ssp585 and ssp126 future projections are set according to the homonymous scenario of socio-economic development, while the rest of simulations include external forcings that are consistent with observations since 1850.

name	period	ensemble	forcings				
		\mathbf{size}	SST	GHG	ozone	aerosols	natural
amip-hist	1870-2014	20	1	1	1	1	1
historical	1850-2014	32	—	1	1	✓	1
hist-GHG	1850-2014	10	-	\checkmark	-	_	-
hist-O3	1850-2014	10	_	—	1	_	_
hist-aer	1850-2014	10	-	_	-	1	-
hist-nat	1850-2014	10	_	_	-	_	1
piControl	1200 years	1	_	_	-	_	-
ssp585 ssp126	2015-2100	6	_	✓*	✓*	✓*	✓*

Table 2 List of observations and reanalyses used in our analyses.

data base	data type	$\begin{array}{c} \textbf{resolution} \\ (^{\circ}lat \times ^{\circ}lon \times \\ vertical \ levels) \end{array}$	period	reference
CR2	rain gauge	in situ	1960-2017	http://www.cr2.cl
GPCCv2018	griddod	$0.5^\circ imes0.5^\circ$	1891 - 2016	Schneider et al. (2018)
UDEL5.01	gridded		1900-2017	Willmott and Matsuura (200
HadISST1	gridded	$1^{\circ} \times 1^{\circ}$	1870-2015	Rayner et al. (2003)
HadSLP2	gridded	$5^{\circ} \times 5^{\circ}$	1850-2017	Allan and Ansell (2006)
NOAA-CIRES 20CR-V2 NOAA-CIRES 20CR-V2c		$2^{\circ} \times 2^{\circ} \times 24$ $2^{\circ} \times 2^{\circ} \times 24$	$\frac{1871-2010}{1851-2014}$	Compo et al. (2011)
NOAA-CIRES 20CR-V3		$1^{\circ} \times 1^{\circ} \times 64$	1836-2015	Slivinski et al. (2019)
ERA-20C	roopolygig	$2^{\circ} \times 2^{\circ} \times 37$	1900-2010	Poli et al. (2016)
ERA-40	reanalysis	$1.1^{\circ} \times 1.1^{\circ} \times 23$	1970-2001	Uppala et al. (2005)
ERA-Interim		$0.75^{\circ} \times 0.75^{\circ} \times 37$	1979 - 2017	Dee et al. (2014)
NCEP1 NCEP2		$\begin{array}{c} 2.5^{\circ} \times 2.5^{\circ} \times 17 \\ 2.5^{\circ} \times 2.5^{\circ} \times 17 \end{array}$	1948-2017 1979-2017	Kalnay et al. (1996)
	CR2 GPCCv2018 UDEL5.01 HadISST1 HadSLP2 NOAA-CIRES 20CR-V2 NOAA-CIRES 20CR-V2 NOAA-CIRES 20CR-V3 ERA-20C ERA-40 ERA-40 ERA-Interim NCEP1	CR2 GPCCv2018 UDEL5.01rain gauge griddedHadISST1griddedHadSLP2griddedNOAA-CIRES 20CR-V2 NOAA-CIRES 20CR-V3 ERA-20C ERA-40reanalysis eanalysisERA-40 NCEP1reanalysis	$\begin{array}{llllllllllllllllllllllllllllllllllll$	$ \begin{array}{c c c c c c c c c c c c c c c c c c c $

Table 3 Ensemble-mean 1970-2014 linear trend in mm decade⁻¹ of the amip-hist precipitation (δ) and of the precipitation anomalies coherent with the corresponding PC1 index (δ_{PC1}) averaged over the boxed areas in Fig. 5. All values are significant with 95% confidence interval.

	MAM	JJA	SON	DJF
δ	-8.54	-6.18	-6.31	-15.15
δ_{PC1}	-8.04	-5.45	-5.17	-15.81

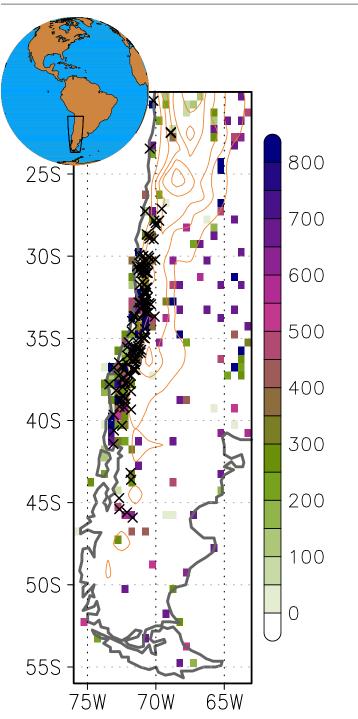


Fig. 1 Details of in situ precipitation observations in South Western South America (SWSA). Crosses: location of the 129 rain gauge stations of the observations used. Colored boxes: number of observed rainfall monthly data accumulated over 1960-2014 at 0.5° horizontal resolution in GPCCv2018. Orange contours: altitude levels in 1000-meter intervals.

Title Suppressed Due to Excessive Length

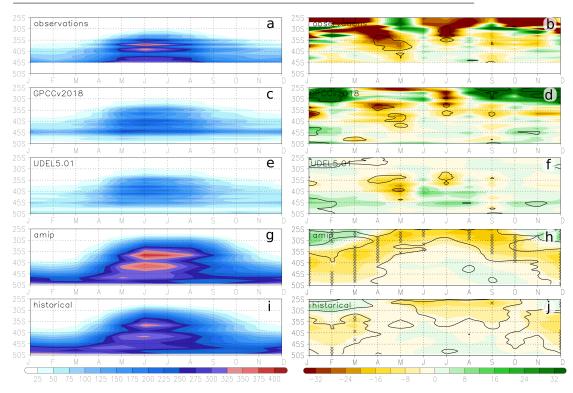


Fig. 2 Hövmoller diagrams representing the climatological annual cycle of monthly precipitation averaged over 1960-2014 (left panel; units are mm) and the linear trend over 1979-2014 (right panel; units are % decade⁻¹) averaged across 75° -70° W from (a-b) the rain gauge observations linearly interpolated to a regular grid of 1° latitude, (c-d) the GPCCv2018 and (e-f) UDEL5.01 gridded observations, (g-h) from the ensemble-mean amip-hist and (i-j) historical simulations. Crosses in (h) and (j) indicate where the trend in 80% of the simulation members show the same sign as the ensemble-mean. Contours in right panel indicate the trend significance at the 5% level according to a Student t-test with number of years minus one degrees of freedom.

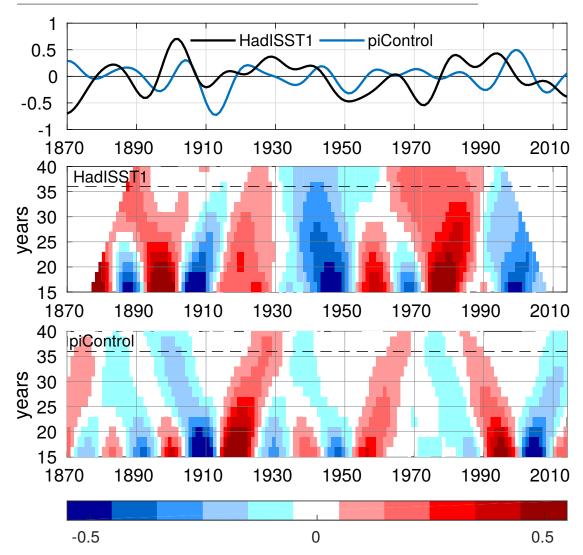


Fig. 3 Observed and simulated amplitude and persistence of the Inter decadal Pacific Oscillation Index (IPO) phase shifts. Upper panel: IPO indices from (black) HadISST1 observations over 1870-2014 and (blue) the piControl run over an equivalent 145-year period. Middle (bottom) panel: Trend values in units of °C per decade of the HadISST1 (piControl) IPO index in centered running windows of varying periods of 15-to-40 years (y axis). Dashed horizontal line indicates the 36-year running window equivalent to the 1979-2014 period.

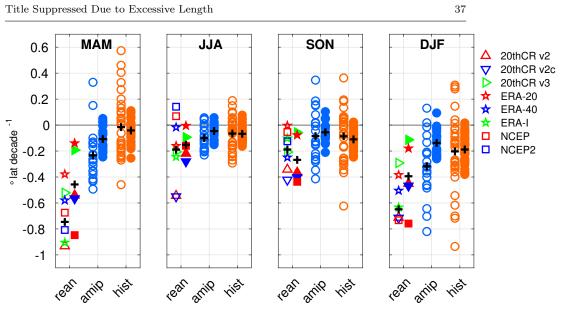


Fig. 4 Linear trend of the seasonal HCE index, expressed in latitude degrees per decade, obtained with reanalyses (rean; left column) and all members of the amip-hist (amip; blue circles) and the historical (hist; orange circles) simulations. Empty shapes correspond to the trend along the common period among all data used (1979-2001) and filled ones to a longer period of 40 years (1971-2010). Black crosses indicate the trend value averaged across all reanalyses or ensemble members.

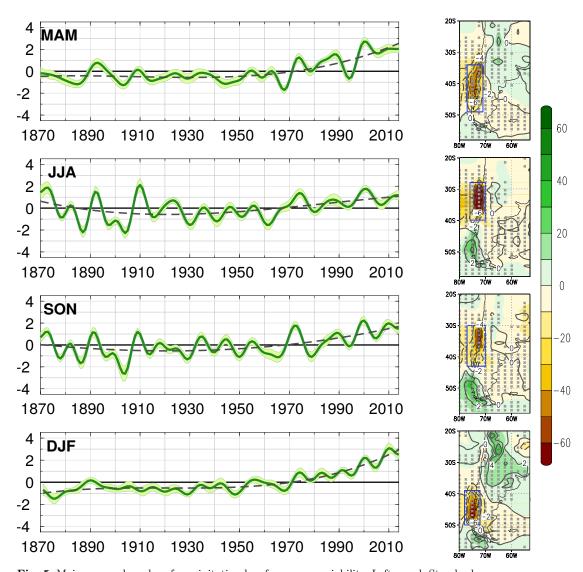


Fig. 5 Main seasonal modes of precipitation low-frequency variability. Left panel: Standardized PC1s of the 8-yr LPF seasonal precipitation anomalies over southern South America (58° - 25° S; 80° - 65° W) from the amip-hist ensemble-mean. Explained variance of the PC1s: 59.1% in MAM, 75.7% in JJA, 67.0% in SON and 69.9% in DJF. Dotted lines indicate the $3^{\rm th}$ -degree polynomial fit and green shading the spread among members at the 95% confidence level. Right panel: Regression patterns of the seasonal precipitation anomalies on the respective seasonal PC1 indices (shading; units are mm per standard deviation) and linear trend (contours; units are mm decade⁻¹) over the entire simulated period, 1870-2014. Gray dots indicate where there is at least 80% agreement among members regarding the sign of the regression coefficient.

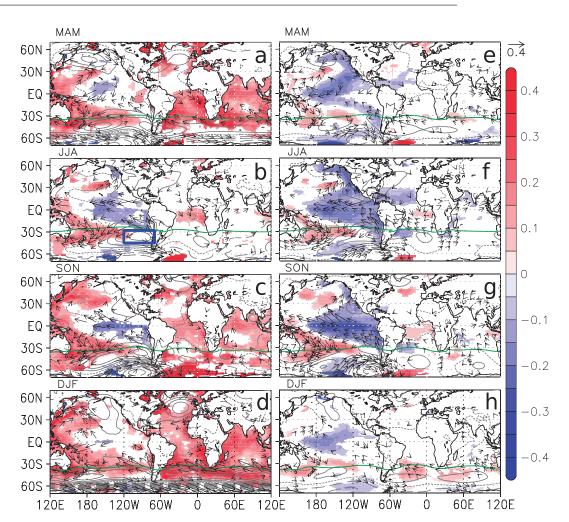
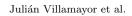


Fig. 6 Seasonal regression patterns of the non-detrended (left) and detrended (right) indices of precipitation on the seasonal anomalies of SST (shaded; units are °C per standard deviation), surface wind (vectors; m s⁻¹ per standard deviation) and SLP (contours in intervals of 0.2 hPa per standard deviation). Regression values of the SST and wind anomalies with lower than 80% statistical significance, according to a random-phase test of Ebisuzaki, are masked. The green lines indicate the position of the climatological maximum of SLP. The patterns are computed using amip-hist ensemble-means. Blue box in panel b) indicate a center of maximum SLP anomalies between 70° - 120° W and 28° - 45° S.



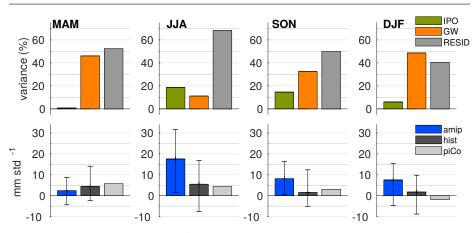


Fig. 7 Decomposition of the SWSA precipitation variance in response to the IPO and GW indices, and simulated precipitation sensitivity to the IPO. Upper panel: Bar charts of the components of the total variance (in %) of the LPF indices of amip-hist ensemble-mean seasonal SWSA precipitation explained by the IPO (green) and GW (orange) indices and the residual (gray) resulting from a multi-linear regression analysis. Bottom panel: Bar charts of the ensemble-mean regression coefficient between the SWSA precipitation seasonal anomalies and the standardized IPO index in amip-hist (blue), historical (black) and piControl (grey) simulations. Error bars indicate the total spread of the result obtained from individual members.

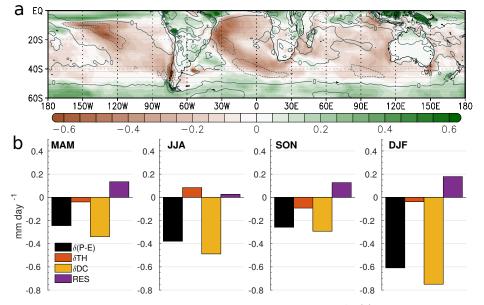


Fig. 8 Decomposition of the net moisture budget change in SWSA. (a) Colors indicate the change in the annual-mean P-E averaged over 2005-2014 with respect to 1851-1910 in units of mm day⁻¹. Contours (in intervals of 0.1 from -0.2 to 0.2 mm day⁻¹ per standard deviation) indicate the regression of the annual P-E anomalies on the standardized annual HCE index over 1850-2014. (b) Bar charts of the seasonal-mean change in the net moisture budget (black), the thermodynamic (orange) and dynamic (yellow) components and the residual (purple) averaged over the corresponding seasonal SWSA areas (see boxes in Fig. 5). All plots are based on the ensemble-mean of the historical simulations.

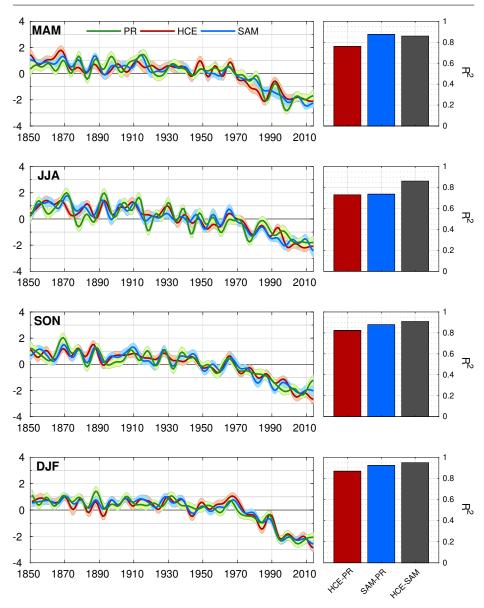


Fig. 9 Left panel: Standardized seasonal indices of (green line) SWSA precipitation (PR), (red line) HCE and (blue line) -SAM from the historical ensemble mean (units are standard deviation). Shading indicates the 95% confidence interval. Right panel: R^2 values between the ensemble-mean (red bar) PR and HCE indices, (blue bar) PR and SAM indices and (gray bar) HCE and SAM indices.

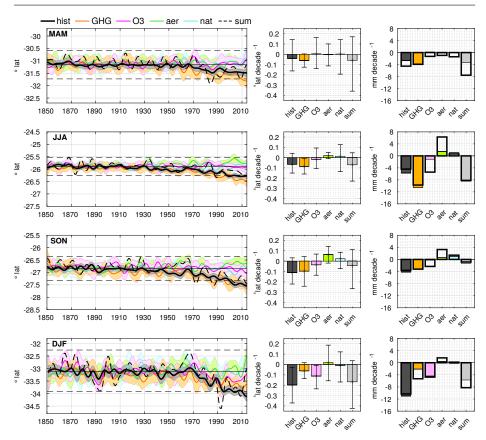


Fig. 10 Left panel: Ensemble-mean HCE indices with 95% confidence interval of the (black) historical, (orange) hist-GHG, (pink) hist-O3, (green) hist-aer, (cyan) hist-nat simulations and the (dashed black) index obtained from the sum of the individual attribution run anomalies. Horizontal lines represent the (solid) climatology and (dashed) 90% dispersion of the piControl HCE anomalies. Middle panel: (color bars) Ensemble-mean HCE index linear trend over 1970-2014 and (error bars) spread from individual members. Right panel: (outlined black bars) Linear trend of the low-frequency index of ensemble-mean HCE index ($\alpha \cdot \delta HCE$; where α is the regression coefficient calibrated over the entire simulated period and δHCE the HCE trend over 1970-2014).

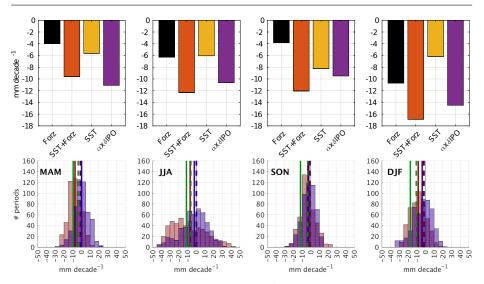


Fig. 11 Upper panel: Trend of the seasonal SWSA precipitation over 1979-2014 induced by external forcing (Forz) (i.e., computed from the historical ensemble-mean), by observed SST and external forcing (SST+Forz) (i.e., computed from the amip-hist ensemble-mean), by only observed SST (i.e., computed as the trend from the amip-hist ensemble-mean minus the one from the historical ensemble-mean) and the trend associated with the trend of the IPO index with a linear regression over 1979-2014 ($\alpha \cdot \delta IPO$). Bottom panel: PDFs of the seasonal SWSA precipitation linear trend (in mm decade⁻¹) over 36-year periods since 1870 in the amip-hist individual members in which the IPO index shows positive (red bars) and negative (blue bars) phase shifts. The red/blue vertical solid lines indicate the trend averaged over all values corresponding to positive/negative IPO shifts. Vertical green lines indicate the 1979-2014 trends in the ensemble-mean amip-hist. Vertical red, blue and green dashed lines represent equivalent trends to the respective colored solid lines but corresponding to the historical simulations.

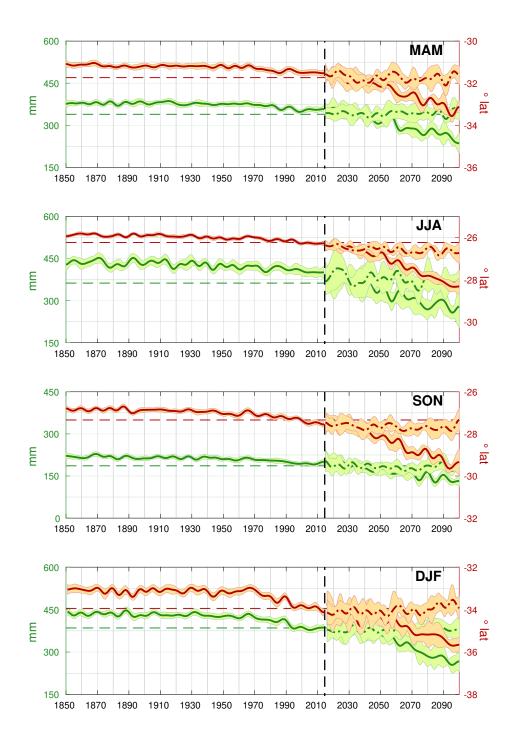


Fig. 12 Seasonal indices of (green) SWSA precipitation and (red) HCE, computed from the ensemble mean of the historical simulations in 1850-2014 and in 2015-2100 of (solid line) the ssp585 and the (dashed line) the ssp126 future projections. Shading indicates the 95% confidence interval. Horizontal dashed lines indicate the lower threshold of the dispersion of the piControl (red) HCE and (green) SWSA precipitation anomalies with 90% confidence interval.

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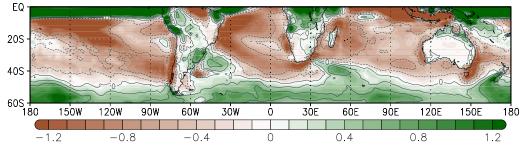


Fig. 13 Projected change of the net moisture budget in the Southern Hemisphere. Colors indicate the change in the annual-mean P-E (units are mm day⁻¹) averaged over 2091-2100 with respect to 2005-2014, based on ensemble-mean ssp585 and historical outputs, respectively. Contours (in intervals of 0.1 from -0.2 to 0.2 mm day⁻¹ per standard deviation) indicate the regression of annual P-E anomalies on the standardized annual HCE index over 2015-2100.

Acknowledgements The authors thank editor Dr. Corti and the two anonymous reviewers 845 for their fundamental duty in the publishing process of this paper and their helpful comments. 846 This work was undertaken in the framework of the THEMES project under the BNP-Paribas 847 Foundation grant and the IPSL Climate Graduate School EUR. The CMIP6 project at IPSL 848 used the HPC resources of TGCC under the allocations 2016-A0030107732, 2017-R0040110492 849 and 2018-R0040110492 (project gencmip6) provided by GENCI (Grand Equipement National 850 de Calcul Intensif). This study benefited from the ESPRI (Ensemble de Services Pour la 851 Recherche l'IPSL) computing and data centre (https://mesocentre.ipsl.fr) which is supported 852 by CNRS, Sorbonne Universite, Ecole Polytechnique and CNES and through national and 853 854 international grants. Special thanks to Juan Rivera for providing us with observational precipitation data from rain gauges and his helpful comments on the model validation. 855

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