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Demographics of Exoplanets in Binaries. I. Architecture of S-type Planetary Systems Revealed by the Radial-velocity Sample

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Abstract

Although the sample of exoplanets in binaries has been greatly expanded, the sample heterogeneity and observational bias are obstacles toward a clear figure of exoplanet demographics in the binary environment. To overcome the obstacles, we conduct a statistical study that focuses on S-type (circumstellar) planetary systems detected by the radial-velocity (RV) method. We try to account for observational biases by estimating, from available RV data, planet detection efficiencies for each individual system. Our main results are as follows. (1) Single (resp. multiple) planetary systems are mostly found in close (wide) binaries with separation $a_B < (>) \sim 100\text{--}300$ au. (2) In binaries, single and multiple-planet systems are similar in 1D distributions of mass and period as well as eccentricity (in contrast to the “eccentricity dichotomy” found in single star systems) but different in the 2D period-mass diagram. Specifically, there is a rectangular-shaped gap in the period-mass diagram of single-planet systems but not for multiples. This gap also depends on binary separation and is more prominent in close binaries. (3) There is a rising upper envelope in the period-mass diagram for planets in wide binaries as well as in single stars but not in close binaries. More specifically, there is a population of massive short-period planets in close binaries but almost absent in wide binaries or single stars. We suggest that enhanced planetary migration, collision and/or ejection in close binaries could be the potential underlying explanation for these three features.

Unified Astronomy Thesaurus concepts: [Exoplanets \(498\)](#)

1. Introduction

Stars are thought to be commonly born and found in binary/multiple systems (Duquennoy & Mayor 1991; Raghavan et al. 2010). Therefore, the demographics of exoplanets in binaries play a crucial role in statistically studying the whole exoplanet population in our Galaxy. Furthermore, the diverse orbital configurations and rich dynamics in planet-bearing binary systems provide valuable conditions to test various theories and models of planet formation and evolution (Thebault & Haghighipour 2015; Marzari & Thebault 2019).

There are generally two orbital configurations of planets in binaries (Dvorak 1982). One is called a P type, where planets orbit both the binary stars, i.e., circumbinary planets like Kepler 413 ABb (Kostov et al. 2013). The other is called an S type, where planets orbit one of the binary stars, i.e., circumstellar planets like γ Cephei Ab (Campbell et al. 1988; Hatzes et al. 2003). Most planets in binaries are currently found in the S type. At the time of writing this paper (2021 March 4), by combining the two catalogs (see Section 2.1 for details) retrieved from the Extrasolar Planets Encyclopaedia and the Catalogue of Exoplanets in Binary Star Systems (Schwarz et al. 2016), there are 211 confirmed S-type planets. Among them, 116 are from radial-velocity (RV) surveys (e.g., γ Cephei Ab Campbell et al. 1988; Hatzes et al. 2003, Tau Bootis b Butler et al. 1997), 89 from transit surveys (e.g., Kepler-68 b Gilliland et al. 2013, HD 202772 b Wang et al. 2019) and six from others, e.g., microlensing surveys (Bennett et al. 2016; Gould et al. 2014) and direct imaging surveys (e.g., 51 Eri Montet et al. 2015). There are only 28 confirmed P-type planets, and a large fraction of them are from transit surveys, e.g., Kepler-16AB b (Doyle et al. 2011) and Kepler-47 b, c, and d

(Orosz et al. 2012, 2019) etc. Note, the above numbers only reflect a small but relatively well characterized portion of planet-bearing binary systems. Many candidate planet-bearing binary systems, e.g., those from imaging followup observations of Kepler planet-host stars (Wang et al. 2014a, 2014b; Kraus et al. 2016; Ziegler et al. 2017; Furlan et al. 2017; Ziegler et al. 2018) are currently not included in the above two catalogs because either the planetary nature or stellar binarity remains to be confirmed.

When mining the observational data, there are two commonly used approaches to study the effect of stellar binarity on planetary systems. On the one hand, by calculating the binary fraction of planet hosts and comparing it to those of nonplanet hosts or field stars, one can learn whether/how planet formation is suppressed by binary stars (Bonavita & Desidera 2007, 2020; Eggenberger et al. 2011; Wang et al. 2014a, 2014b; Kraus et al. 2016; Ziegler et al. 2017, 2018; Matson et al. 2018; Moe & Kratter 2019; Fontanive et al. 2019; Fontanive & Bardalez Gagliuffi 2021). One difficulty with this approach is to remove or at least quantify the effects of observational bias against finding planets in binaries. A major bias often comes from the target selection process itself, which is severely biased against binaries in most RV surveys; but also significant for transit surveys, e.g., Kepler (Wolniewicz et al. 2021). On the other hand, by investigating the distributions of planet properties (e.g., mass, period, etc.) in known planet-bearing binary systems, one can learn how planetary architecture is sculpted by binary stars, e.g., (Roell et al. 2012). It is the latter approach that we adopt in this paper. Furthermore, we only focus on S-type planet systems from RV surveys to reduce the sample complexity and thus better take into potential observational bias. Since we just care about “relative” properties (e.g., period-mass distribution) rather than “absolute”

properties (e.g., absolute occurrence rate), our analysis is not affected by binary-hostile target selection biases in RV surveys. The main potential bias we need to correct for is the varying planet detectability performances depending on planetary mass and period, which could overrepresent (or underrepresent) some types of planets in our sample. We could also have, for the same location in the planetary mass-period diagram, different detectability thresholds depending on the considered system. For example, some systems may have poorer RV measurements due to large stellar activity or light contamination of close-binary stars, which would thus bias against finding smaller planets on longer periods in these systems.

Some previous statistical studies have already explored the properties of S-type planets in RV surveys, but for more limited samples. With a small sample of 9 S-type planets, Zucker & Mazeh (2002) found that planets in binaries seem to follow a different period-mass distribution than planets in single star systems. Increasing the sample size to 19, Eggenberger et al. (2004) confirmed this trend by pointing out that the few most massive ($M_{\text{Sini}} > 2M_{\text{Jup}}$) short-period ($P < 40$ days) planets are exclusively in binary systems. Doubling the sample size to ~ 40 , Desidera & Barbieri (2007) found that the above mass difference is significant only in tight binaries (binary separation, $a_B < 100 - 300$ au). Duchêne (2010) then found that planets in tight binaries ($a_B \lesssim 100$ au) are significantly more massive than those in wide binaries ($a_B \gtrsim 100$ au). Roell et al. (2012) compiled a larger sample of planets in binaries and found that (1) no planets are in binaries with a projected separation of less than 10 au; (2) multiplanets systems are exclusively in wide binaries ($a_B > 100$ au); and (3) planetary mass decreases with an increasing binary separation.

In this paper, which is the first (Paper I) of a series of studies on the Demographics of Exoplanets in Binaries, we focus on the architecture of S-type planetary systems detected by the RV method. With the significantly expanded sample that we compile (110 S-type RV planets), we aim to (1) finding new patterns and (2) revisiting the above reported patterns—e.g., the period-mass distribution—by taking into account the RV detection efficiency in detail.

This paper is organized as the follows. In Section 2, we describe our planet sample and the method used to consider the effect of the RV detection efficiency. In Section 3, we present our results. Finally, we discuss and summarize the paper in Sections 4 and 5.

2. Sample and Method

2.1. Sample

We build our RV planet sample by retrieving data (on 2021 March 4) from two online catalogs: (1) the catalog of exoplanet-hosting binaries with separations up to 500 au;³ (2) the catalog of exoplanets in binary star systems (Schwarz et al. 2016).⁴ Specifically, we first retrieved all RV planet systems in the first catalog, which is almost exhaustive for all binaries of separation less than ~ 500 au. Then, we turned to the second catalog, which provided most planets in binaries with separation from ~ 500 up to $\sim 15,000$ au in our sample. We note that HD 7449c is reported as planet with minimum mass of $19 m_J$ in the second catalog but not recognized in the first catalog. In fact, HD

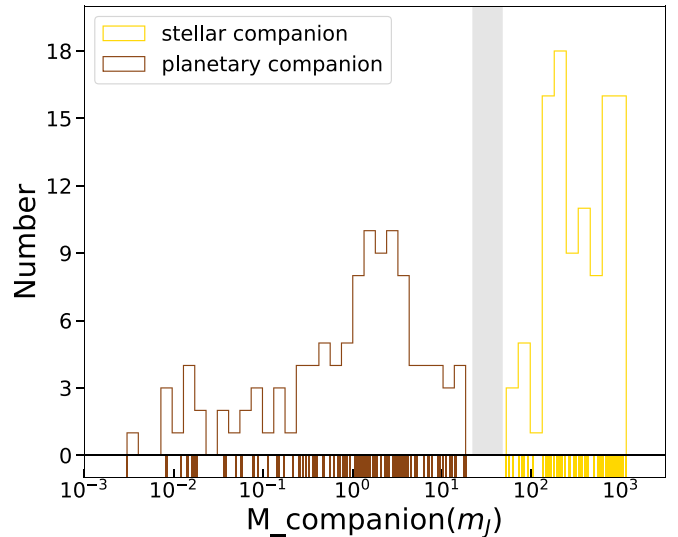


Figure 1. The mass (or M_{Sini}) histogram of all companion objects in our sample. Apparently, there is gap (gray shaded) between 20 and 50 Jupiter masses, which separates planetary objects from stellar objects.

7449c is only inferred from the long-term RV trend that could be due to the stellar companion (Rodigas et al. 2016). Therefore, we removed this unconfirmed planet and thus HD 7449 is included as a single-planet system in our sample.

Furthermore, in order to have a unified criterion to define our sample, we adopted an upper mass limit ($20 m_J$) for planetary objects and a lower-mass limit ($50 m_J$) for stellar objects. This criterion is motivated by the mass (or M_{Sini}) distribution of the companion objects in Figure 1, which shows a gap between 20 and 50 Jupiter masses. Under this definition, HD 87646c and HD 4113c were treated as stellar objects instead in our sample.

In addition, systems HD 41004 AB, HD 20782/HD 20781, and HD 133131 AB have RV planets orbiting both the primary and companion stars, and they are treated as two single S-type planet systems in our analyses. Systems HD 80606, HD 20781, HD 103774, and HD 93385, whose RV data are not publicly available, are thus excluded from our analyses.

With these criteria, we obtain 110 planets in 80 systems, 61 (19) of which are single (multiple) planet systems. Table 1 lists the properties of these planets in our sample. Figure 2 is an overview of our sample in the plane of $M_{\text{Sini}}-a_B$, where M_{Sini} is the minimum mass of planet and a_B is the semimajor axis of binary orbit. Our samples also include 23 planets in 16 triple star systems and quadruple star systems. In these cases, a_B reported in Table 1 should be treated as the semimajor axis of the closest planet-host companion.

2.2. Quantifying the RV Detection Constraints

Different systems could differ significantly in the precision of their RV measurements, causing an observational bias that systems with poorer RVs bias against finding smaller planets with longer period. In order to quantify this bias, we perform the following RV-signal injection-recovery test in each system of our sample.

First, we collect the RV data of each system and extract the RV residual, RV_{residual} , by subtracting the observed planet signals using the RV-fitting tool *Systemic2* (<http://www.stefanom.org/console-2/>; Meschiari et al. 2009). During the above RV-fitting process, we also removed the long-term RV trend from binary orbits in some systems (e.g., Gliese 15,

³ http://exoplanet.eu/planets_binary/

⁴ <https://www.univie.ac.at/adg/schwarz/multiple.html>

Table 1
S-type Planet in Binary Stellar Systems

Row_Id (1)	Planet_id (2)	a_B (au) (3)	a_c (au) (4)	N_p (5)	$M_p(M_J)$ (6)	a_p (au) (7)	Period(days) (8)	Eccentricity (9)	Source (10)	N-star (11)
1	HD 42936 b	1.22	0.2	1	0.0081	0.0662	6.6732	0.14	1	2
2	HD 87646 b	1.58	0.2	1	12.4	0.117	13.481	0.05	1	2
3	HD 59686 A b	13.56	1.0	1	6.92	1.086	299.36	0.05	1	2
4	HD 7449 b	18	2.9	1	1.09	2.33	1270.5	0.92	1	2
5	gamma Cephei b	19	3.7	1	1.85	2.05	903.3	0.049	1	2
6	HD 4113 Ab	23.7	4.6	1	1.56	1.28	526.62	0.903	1	2
7	HD 41004 Ab	26.48	5.5	1	2.54	1.64	963	0.39	1	2
8	HD 41004 Bb	26.48	5.5	1	18.4	0.0177	1.3236	0.058	1	2
9	30 Ari b	27.6	1.7	1	13.82	0.99	335.1	0.29	1	4
10	GJ 86 b	27.8	7.8	1	4.01	0.11	15.766	0.046	1	2
11	HD 196885 b	29	6.5	1	2.98	2	1326	0.48	1	2
12	HD 2638 b	32.2816	6.8	1	0.48	0.044	3.438	0.041	1	3
13	HD 8673 b	35	5.4	1	14.2	3.02	1634	0.723	1	2
14	HD 126614 b	45.6482	10.9	1	0.38	2.35	1244	0.41	1	3
15	HD 164509 b	46.0265	10.5	1	0.48	0.875	282.4	0.26	1	2
16	GJ 3021 b	85.747	22.3	1	3.37	0.49	133.71	0.511	1	2
17	Gliese 15 b	93	12.7	2	0.017	0.072	11.44	0.27	1	2
18	Gliese 15 c	93	12.7	2	0.114	5.4	7391	0.27	1	2
19	HD 30856 b	117.273	25.6	1	1.8	2	912	0.24	1	2
20	HD 177830 b	123.57	31.9	2	1.49	1.2218	406.6	0.001	1	2
21	HD 177830c	123.57	31.9	2	0.15	0.5137	110.9	0.35	1	2
22	HD 142 b	132.405	28.5	2	1.25	1.02	349.7	0.17	1	2
23	HD 142 c	132.405	28.5	2	5.3	6.8	6005	0.21	1	2
24	HD 79211	136.86894	27.3	1	0.3231	0.141	24.45	0.11	1	2
25	HD 207832 b	142.493	35.3	2	0.56	0.586	160.07	0.197	1	4
26	HD 207832c	142.493	35.3	2	0.73	2.112	1155.7	0.27	1	4
27	HD 197037A b	150.059	35.4	1	0.79	2.07	1035.7	0.22	1	2
28	HD 195019 b	165.19	33.0	1	3.7	0.1388	18.2016	0.014	1	2
29	HD 114762 b	166.452	42.8	1	10.98	0.353	83.9151	0.3354	1	2
30	HD 217786 b	197.97	51.1	1	13	2.38	1319	0.4	1	2
31	HD 98736 b	199.23	39.7	1	2.33	1.864	968.8	0.226	1	2
32	HD 116029 b	215.63	55.4	2	2.1	1.73	670.2	0.21	1	2
33	HD 116029c	215.63	55.4	2	1.27	2.144	907	0.038	1	2
34	HD 19994 b	220	57.2	1	1.68	1.42	535.7	0.3	1	2
35	tau Bootis b	221	5.3	1	4.13	0.046	3.3135	0.0787	1	2
36	HD 212301 b	238	57.2	1	0.4	0.036	2.2457	0.0147	1	2
37	HD 142245A b	244	50.5	1	1.9	2.77	1299	0.32	1	3
38	HD 65216 b	256.8	68.4	1	1.21	1.37	613.1	0.41	1	3
39	HD 185269A b	271.11	69.0	1	0.94	0.077	6.838	0.3	1	2
40	HD 16141 b	281.203	67.2	1	0.215	0.35	75.82	0.28	1	2
41	Gliese 667 b	300	54.1	6	0.0176	0.0505	7.2	0.13	1	3
42	Gliese 667 c	300	54.1	6	0.0119	0.125	28.14	0.02	1	3
43	Gliese 667 d	300	54.1	6	0.016	0.276	91.61	0.03	1	3
44	Gliese 667 e	300	54.1	6	0.0085	0.213	62.24	0.02	1	3
45	Gliese 667 f	300	54.1	6	0.0085	0.156	39.03	0.03	1	3
46	Gliese 667 g	300	54.1	6	0.0145	0.549	256.2	0.08	1	3
47	HD 27442 b	302.64	61.4	1	1.35	1.16	415.2	0.058	1	2
48	HD 28254A b	302.64	60.3	1	1.16	2.15	1116	0.81	1	2
49	HD 106515 b	345	49.1	1	9.61	4.59	3630	0.572	1	2
50	HD 86081A b	348	95.0	1	1.5	0.039	1.99809	0.0575	1	2
51	HD 114729 b	355.602	85.6	1	0.84	2.08	1135	0.32	1	2
52	HD 133131 Ab	379.22	67.6	2	1.43	1.44	649	0.32	1	2
53	HD 133131 Ac	379.22	67.6	2	0.48	4.36	3407	0.47	1	2
54	HD 133131 Bb	379.22	66.8	1	2.5	6.415	6119	0.62	1	2
55	HD 108341 b	383	76.4	1	3.5	2	1129	0.85	1	2
56	HD 102365 b	406.042	99.9	1	0.05	0.46	122.1	0.34	1	2
57	HD 220842A b	428.739	85.5	1	3.18	0.74	218.47	0.404	1	2
58	HD 132563 b	430.75	74.7	1	1.49	2.62	1544	0.22	1	3
59	HD 46375 b	435	87.8	1	0.249	0.041	3.024	0.0524	1	2
60	HD 43691 b	452.69	111.1	1	2.57	0.238	36.99913	0.085	1	3
61	omi UMaA b	518.271	103.3	1	4.1	3.9	1630	0.191	1	2
62	HD 180617 b	560.6406	144.8	1	0.0384	0.3357	105.9	0.16	1	2
63	HD 4732c	605.9105	120.8	2	2.37	4.6	2732	0.23	1	2
64	HD 4732 b	605.9105	120.8	2	2.37	1.19	360.2	0.13	1	2
65	HD 109749 b	618.5205	121.3	1	0.28	0.0635	5.24	0.01	2	2
66	HD 196050 b	631.761	154.2	1	2.83	2.47	1316.24	0.21	2	3
67	HD 178911 b	640	112.4	1	6.29	0.32	71.49	0.124	2	3
68	HD 96167 b	645.632	172.2	1	0.68	1.3	498.9	0.71	2	2
69	HD 99492 b	649.415	121.7	2	0.087	0.122	17.1668	0.13	2	2
70	HD 99492c	649.415	121.7	2	0.359	5.4	13.1	0.1	2	2
71	Ups And b	750	194.5	4	0.62	0.059	4.6171	0.013	2	2

Table 1
(Continued)

Row_Id (1)	Planet_id (2)	a_B (au) (3)	a_c (au) (4)	N_p (5)	$M_p(M_J)$ (6)	a_p (au) (7)	Period(days) (8)	Eccentricity (9)	Source (10)	N-star (11)
72	Ups And c	750	194.5	4	1.8	0.861	237.7	0.24	2	2
73	Ups And d	750	194.5	4	10.19	2.55	1302.61	0.274	2	2
74	Ups And e	750	194.5	4	1.059	5.2456	3848.86	0.005	2	2
75	HD 75289 b	781.819	204.8	1	0.47	0.046	3.5098	0.021	2	2
76	HD 188015 b	862	217.6	1	1.26	1.19	456.46	0.15	2	2
77	HD 147379 b	869.4	171.8	1	0.0777	0.3193	86.54	0.01	2	2
78	GJ 676 b	1010	247.8	4	4.95	1.8	1050.3	0.328	2	2
79	GJ 676 c	1010	247.8	4	3	5.2	4400	0.2	2	2
80	GJ 676 d	1010	247.8	4	0.014	0.0413	3.6	0.15	2	2
81	GJ 676 e	1010	247.8	4	0.036	0.187	35.37	0.24	2	2
82	HD 142022 b	1030	210.6	1	5.1	3.03	1928	0.53	2	2
83	55 Cnc b	1065	278.6	5	0.8	0.1134	14.651	0.016	2	2
84	55 Cnc c	1065	278.6	5	0.169	0.2403	44.3446	0.053	2	2
85	55 Cnc d	1065	278.6	5	3.835	5.76	5218	0.025	2	2
86	55 Cnc f	1065	278.6	5	0.144	0.781	260.7	0	2	2
87	16 Cyg b	1080	189.7	1	1.68	1.68	799.5	0.689	2	3
88	HD 11964 b	1272.6	258.8	2	0.622	3.16	1945	0.041	2	2
89	HD 11964c	1272.6	258.8	2	0.079	0.229	37.91	0.3	2	2
90	HD 101930 b	2772	507.2	1	0.3	0.302	70.46	0.11	2	2
91	91 Aqr b	2835	554.8	1	3.2	0.7	181.4	0.03	2	3
92	HD 89744 b	3100	856.3	2	7.2	0.88	256	0.7	2	2
93	HD 89744c	3100	856.3	2	3.2	0.44	85.2	0.29	2	2
94	HD 190360 b	3585.96	905.3	2	1.502	3.92	2891	0.36	2	2
95	HD 190360c	3585.96	905.3	2	0.057	0.128	17.1	0.01	2	2
96	HD 213240 b	4915.378	1299.8	1	4.5	2.03	951	0.45	2	3
97	XO-2S b	5796	1027.0	2	0.259	0.1344	18.157	0.18	2	2
98	XO-2S c	5796	1027.0	2	1.37	0.4756	120.8	0.153	2	2
99	HD 222582 b	5990	1416.5	1	7.75	1.35	572.38	0.725	2	2
100	HD 125612 b	5990	1536.1	3	3	1.37	502	0.46	2	2
101	HD 125612c	5990	1536.1	3	0.058	0.05	4.1547	0.27	2	2
102	HD 125612 d	5990	1536.1	3	7.2	4.2	3008	0.28	2	2
103	HD 147513 b	6760	1326.9	1	1.21	1.32	528.4	0.26	2	2
104	HD 40979 b	8070.4	1764.6	1	3.28	0.83	263.1	0.25	2	3
105	Prox Cen b	10962	2041.7	2	0.003	0.0485	11.186	0	2	3
106	Prox Cen c	10962	2041.7	2	0.0182	1.48	1894	0	2	3
107	HIP 70849 b	11340	3081.9	1	9	10	10000	0.6	2	2
108	HD 20782 b	11550	2386.2	1	1.9	1.381	591.9	0.97	2	2
109	HD 38529c	15184.962	3806.5	2	17.7	3.695	2134.76	0.36	2	2
110	HD 38529 b	15184.962	3806.5	2	0.87	0.131	14.31	0.25	2	2

Note. a_B : the semimajor axis of binary orbit, calculated via the empirical relationship (Fischer & Marcy 1992) $a_B/r = 1.261$ (r is projected binary separation) if there is no reported a_B in the source catalog. a_c : the critical distance to the primary star to maintain long-term orbital stability (Holman & Wiegert 1999). Source: 1, the catalog of exoplanet-hosting binaries with separations up to 500 au; (http://exoplanet.eu/planets_binary/) 2, the catalog of exoplanets in binary star systems (Schwarz et al. 2016). (<https://www.univie.ac.at/adg/schwarz/multiple.html>) N-star: 2, binary star system; 3 or 4, triple system or quadruple system.

gamma Cephei, tau Boot A, HD41004 A/B, HD196885). For a few systems (HD7449, HD 8673, HD59686, HD87646, HD2638, HD 30856, HD 42936, HD126614, and 30 Ari), the RV_{residual} were extracted directly from the figures (converting graph into numbers) in the corresponding planet discovery papers because the RV data were not readily available.

Next, we inject a test RV signal, i.e., RV_{signal} of a hypothetical planet with a given minimum mass (M_{sini}) and orbital period (P). Ignoring the effect of orbital eccentricity, we have

$$RV_{\text{signal}} = \sqrt{\frac{2\pi G}{P} \frac{(M_{\text{sini}})^3}{(M_* + M)^2}} \cos(\omega + \frac{2\pi t}{P}), \quad (1)$$

where M_* is the mass of host star, M is the mass of hypothetical planet, G is gravitational constant, i is the inclination between the normal of the planet's orbit and the line of sight, ω is the argument of periastron (drawn randomly from 0 to 2π) of the planet's orbit, and t is observing time.

Then, we generate a set of simulated RV data,

$$RV_{\text{sim}} = RV_{\text{signal}} + RV_{\text{residual}} + RV_{\text{error}}, \quad (2)$$

by combining the test planet signal with the RV residual and the RV-measurement error. Here, RV_{error} is a random number that follows the normal distribution, i.e., $N(0, \sigma_{\text{RV}})$, where σ_{RV} is the reported RV uncertainty.

Finally, we evaluate whether the injected planet signal can be recovered from the simulated RV data. The criteria are (1) the peak location of the periodogram (Lomb 1976; Scargle 1982) should be within 25% away from the orbital period of the injected planet, i.e., $|P_{\text{peak}} - P_{\text{inject}}| < 0.25 * P_{\text{inject}}$ and (2) the False Alarm Probability (FAP) of the periodogram peak should be less than 1%. The FAP is calculated via a bootstrap method (Townsend 2010). Specifically, we randomly shuffled the simulated RV data from Equation (2), and record the power of the periodogram peak (Lomb 1976; Scargle 1982). We repeat such a random calculation 1000 times, and the FAP

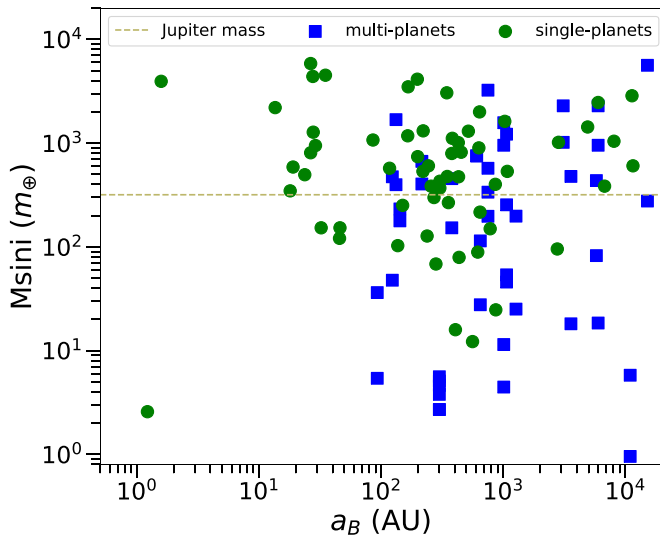


Figure 2. Overview of our sample: planet mass as a function of binary separation. All the single-planet systems are represented by green circles, while all planets in multiple systems are represented by blue squares. The light-yellow dash line indicates the Jupiter mass.

is set as the fraction of times when the record power is larger than the original one.

Figure 3 shows the results of the RV-signal injection-recovery tests for two individual systems (30 Ari and HD 125612; see also in Figures 12–15 in the Appendix for the results of all systems of our sample). For each system, the RV signals are generated by hypothetical planets which are uniformly distributed in the $\log(P)$ – $\log(M_{\text{sini}})$ plane. As can be seen, the filled dark-gray circles show where the hypothetical planets can be recovered from the simulated RV data. The bottom envelopes of the recovered region generally follow the line where $\text{RV}_{\text{signal}} = 1 - 2 \times \text{rms}$, and there is a long period cutoff due to the time baseline of the RV data. On the left is 30 Ari, a close binary ($a_B = 27.6$ au) hosting a single massive planet candidates ($M_{\text{sini}} = 13.82M_J$), while on the right is HD 125612, a wide binary ($a_B = 5990$ au) hosting three planets with masses of $M_{\text{sini}} = 7.2, 3.0, 0.058M_J$. Figure 3 illustrates a potential detection bias of the RV method. Indeed, among the three planets of HD 125612, only the most massive one is close to the recovery region of 30 Ari, while the other two less massive planets would not have been discovered even if they were truly existing in the 30 Ari system. This could be due to the fact that, because RV precision is generally poorer in close binaries (due to light contamination or inappropriate stellar properties e.g., large v_{sini}), only the signal of the largest planets of these systems can be retrieved, leaving potential additional smaller planets undetected (whereas such planets would have been detected in wider binaries or single stars). As a consequence, the fraction of single-planet systems could be artificially overestimated in close binaries. To address this potentially crucial issue, we will have to take into account the statistical average of detection limits for our respective samples of single-planet and multiple-planet systems.

3. Results

3.1. Single Planets versus Multiple Planets

We first compare the properties of single- and multiple-planet systems. Figure 4 shows their period-mass diagrams. In order to

visualize the aforementioned potential RV detection bias, we also overplot the detection limits of each individual system, i.e., the boundaries between the recovery and no-recovery regions in the RV-signal injection-recovery test (e.g., Figure 12). We then estimate the medians of the detection limits for each subsample of single-planet and multiple-planet systems (green and blue curves in Figure 4), which are defined by the 50% completeness contours, meaning that planets below them in the period-mass diagram could not be detected in half systems of the samples. To reduce the effects of such observational biases, we remove planets in the regions with low completeness. Specifically, we implement a cut (hereafter labeled “cut-1”) for which we remove planets below the median detection limit of single-planet systems (the green curve) both in the single and the multiple subsamples. Note that the resulting planetary parameters actually depend on the number of planets considered when fitting the RV data (e.g., Wittenmyer et al. 2013). For simplicity, we ignored this effect in this paper and adopted the same planetary parameters before and after “cut-1”.

In addition, we also consider another cut (hereafter “cut-2”) for which we remove systems with very poor RV data (i.e., $\text{rms} > 100 \text{ m s}^{-1}$), since the detection limits of such poor RV systems deviate too much from the median, and the detectable areas are too small in the period-mass diagram (e.g., system HD 87646; see Figure 12). Figure 4 shows that there are 61 single-planet systems, of which 53 are above the detection limit of cut-1 (the green curve); while there are 53 planets in multiple systems, of which only 26 lie above cut-1. Note that singles and multiples in Figure 4 are initially defined according to the number of planets actually observed in the systems. However, because of the significant fraction of planets that do not pass the cut-1 test, when analyzing statistical properties, we also consider another classification, for which we define single- and multiple-planet systems according to the number of “remaining” planets after applying this cut-1 criterion (see Section 3.1.2). With this alternative classification, some initial multiple-planet systems could be converted into single-planet ones (see Figures 5 and 6).

3.1.1. Multiple-planet Systems Prefer Wide Binaries

Previous studies have hinted at that multiple-planet systems are preferentially found in wide binaries (Roell et al. 2012). Here, we revisit this trend with our larger and updated sample. The three panels in the first column of Figure 5 compare the CDFs of binary separations (a_B) for single- and multiple-planet systems. Already in the raw unbiased sample there is clear tendency for multiple-planet systems to have larger a_B values as compared to that of single-planet systems. This trend is pronounced with a KS test p value of $= 0.00054$, and it is still significant (p value of $= 0.00062$) when attempting to debias the sample by applying cut-1, or with both cut-1 and cut-2 (p value of $= 0.001$). Our results thus clearly confirm that multiple-planet systems are more prevalent in wide binaries whereas single-planet ones are, on average, found in tighter binaries.

In addition, we also consider the dynamical stability limit, a_c , defined as the critical distance to the primary star beyond which orbits become unstable due to the perturbations of the companion star. We estimate a_c using the empirical formula given by Holman & Wiegert (1999), which is a function of binary mass ratio, orbital semimajor axis (a_B) and eccentricity (e_B). We adopted a typical $e_B = 0.3$ (Duquennoy & Mayor 1991) if the binary eccentricity is not available from the source catalog. The right panels of Figure 5 compare the a_c distributions

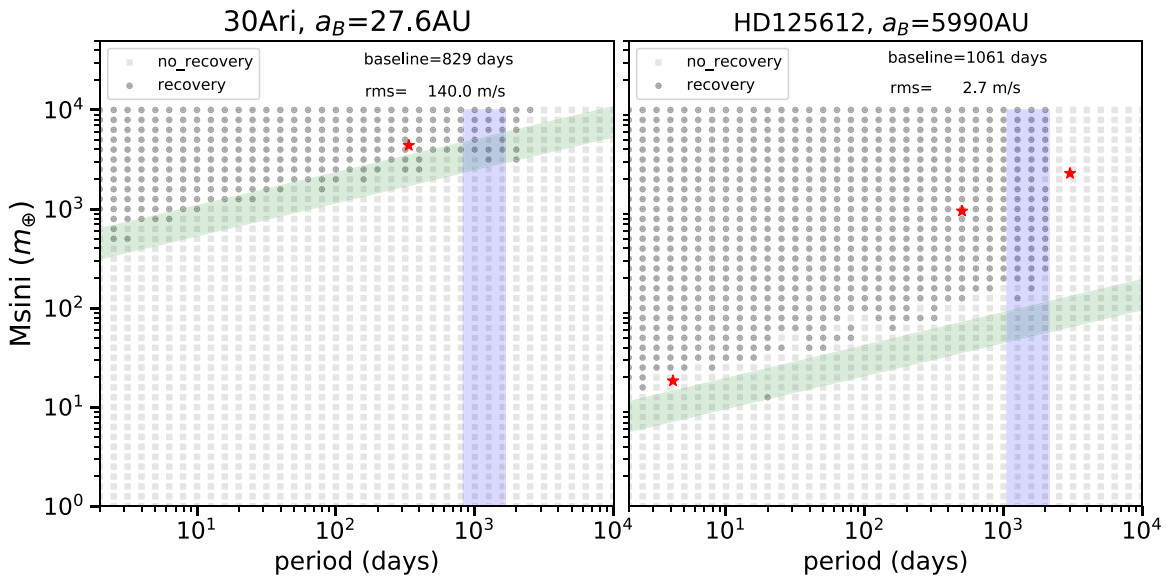


Figure 3. Two examples of RV-signal injection-recovery test. 30 Ari is a close binary system with one planet (red star in the left panel), and HD 125612 is wide binary with three planets (red stars in the right panel). The filled dark-gray circles (light squares) show where the hypothetical planets can (cannot) be recovered from the simulated RVs. The green shaded regions mark where $RV_{\text{signal}} = (1-2) \times rms$. The purple shaded regions mark where $period = (1-2) \times \text{the time baseline}$ of the simulated RV data.

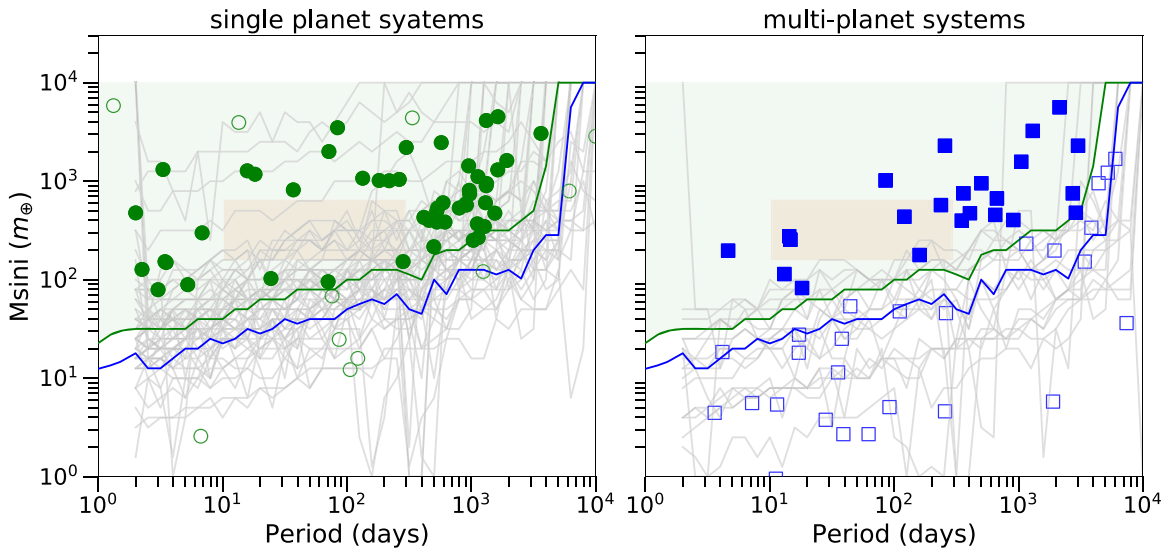


Figure 4. Period-mass diagrams for single-planet systems (green circles in the left panel) and multiple-planet systems (blue squares in the right panel). The green and blue curves represent the medians of the detection limits (gray curves) of singles and multiples, respectively. Systems with RV rms larger than 100 m s^{-1} or planets below the single median detection limit (green curves) are marked with open symbols. The shaded rectangle marked the gap, a planet desert in the singles (left panel; see more in the text in Section 3.1.3).

between single and multiples. We see that multiple planetary systems not only prefer “physically wide” (large a_B) binaries but also “dynamically wide” (large a_c) binaries. This trend is significant with a KS test p value of less than 0.001 regardless of whether cut-1 and cut-2 are applied or not.

3.1.2. Period, Mass, and Eccentricity Distributions

The three columns of Figure 6 compare the distributions of planet properties, i.e., orbital period, minimum mass, and eccentricity for single- and multiple-planet systems. As can be seen, singles and multiples show similar period distributions regardless of whether the cuts correcting for observational bias are applied or not. For mass and eccentricity, singles, as compared to multiples, seem to be more massive (p

value = 0.01) with larger eccentricities (p value = 0.02) with a statistical confidence level about 2-sigma if using the whole “raw” sample. However, this trend disappears when correcting for observational bias by applying cut-1 and cut-2, for which the mass and eccentricity distributions become statistically indistinguishable (KS test p values larger than p value of 0.4) between singles and multiples. As a consequence, the initial trend in the raw data is probably only due to the fact that an important fraction of the smaller planets with lower eccentricities observed in multiple systems could not have been detected in the singles (see Figure 4).

Note that, as mentioned earlier, in the second and third rows of Figure 6, systems are classified as singles/multiples according to the number of planets remaining after correcting for detection

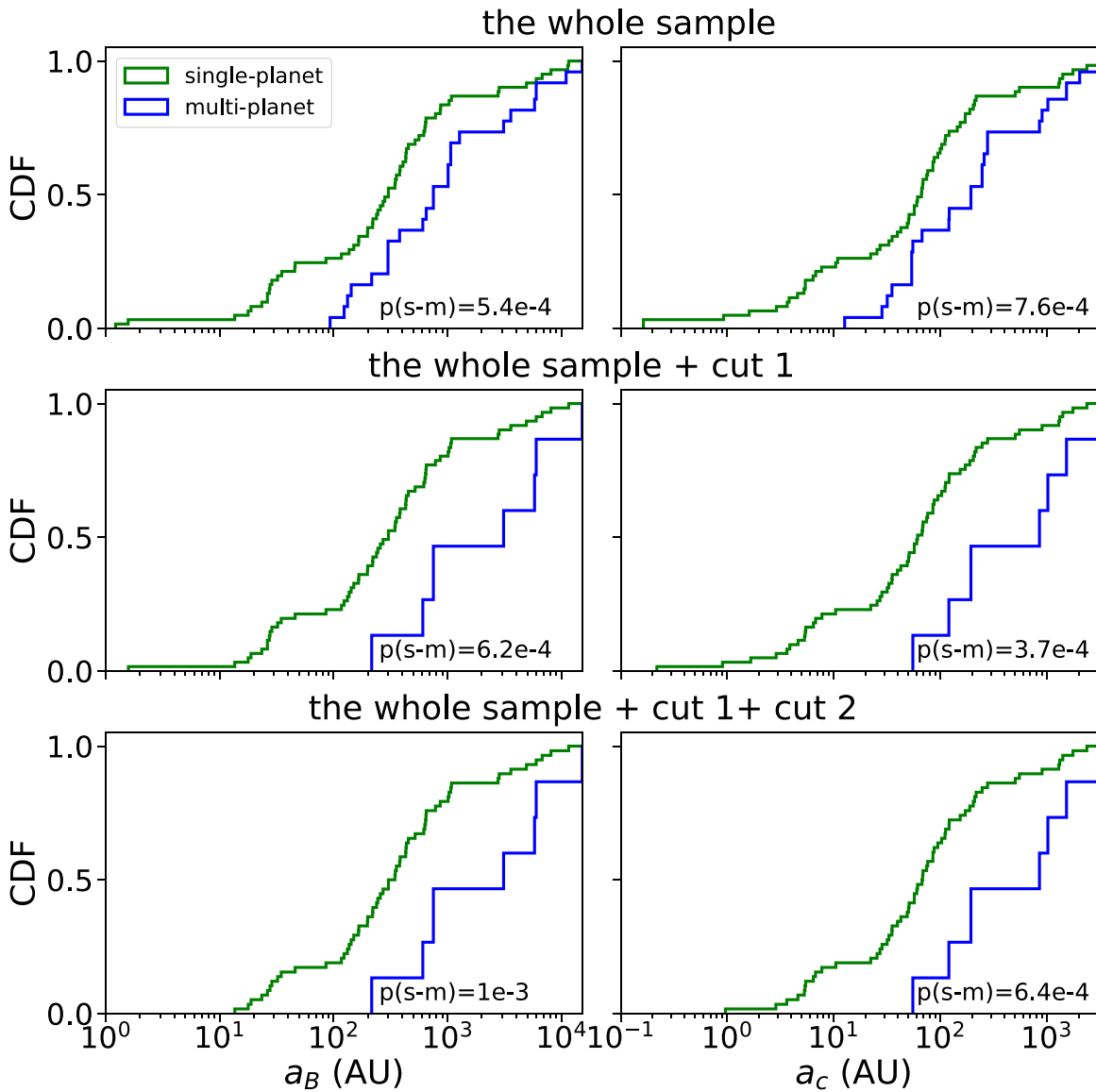


Figure 5. Cumulative distribution functions (CDF) of the semimajor axis of binary orbit (a_B , first column), and the critical distance for long-term orbital stability (a_c , second column). *Top row:* “raw” sample defined without taking into account potential observational biases. *Middle row:* removing planets below the median detection limit of the singles “cut-1”. *Bottom row:* with both cut-1 and cut-2, i.e., removing systems with RV rms > 100 ms^{-1} . Note that, in the middle and bottom rows, the single- and multiple-planet samples are defined according to the number of planets left *after* applying the different detection limit cuts (so that, for example, some “multiple systems” of the top row are categorized as “single-planet” systems in the middle row). In each panel, we perform the two sample KS test and print the p value of the test between singles and multiples, i.e., $p(s-m)$.

limits (cut-1 and cut-2). More specifically, 11 multiple-planet systems from the raw sample are “converted” into singles after applying cut-1 (61 singles and 15 multiples in the second row) and two more if further applying cut-2 (58 singles and 15 multiples in the third row). Nevertheless, we find similar results regarding the statistics of planetary period, mass and eccentricity distributions (no difference between single- and multiple-planet systems) even when keeping the initial criterion for defining the two subsamples (i.e., a “multiple” system is still labeled as such even if it only has one planet surviving cuts 1 and 2), which confirms the robustness of this result.

In summary, for giant planets in binary star systems (all planets that lie above the median detection limit have $M_{\text{sini}} > 60m_{\oplus}$), we find that their 1D distributions in period, mass, and eccentricity is similar in single and multiple-planet systems. As for smaller planets ($M_{\text{sini}} < 60m_{\oplus}$), the conclusion cannot be drawn with current RV data.

3.1.3. A Gap in the Period-mass Diagram of Single-planet Systems?

Although singles and multiples do not show significant differences in the 1D distribution of either period or mass (Figure 6), they seem to differ significantly in the 2D distribution of the period-mass diagram. As can be seen in Figure 4, there are $N_s = 53$ singles above the median detection limit (green curve) but none is located in the middle shaded rectangle region (period = 10–300 days and $M_{\text{sini}} = 160\text{--}640 m_{\oplus} \sim 0.5\text{--}2.0 m_j$). For comparison, there are $N_m = 23$ multiples above the green detection limit curve and 5, a fraction of $f = 0.22$ are located in the middle shaded rectangle. If we adopt a null hypothesis by assuming that singles and multiples follow the same 2D distribution in the period-mass diagram, then we can use binominal distribution to access the significance of the observed gap. The probability that no more than m points are found in the gap

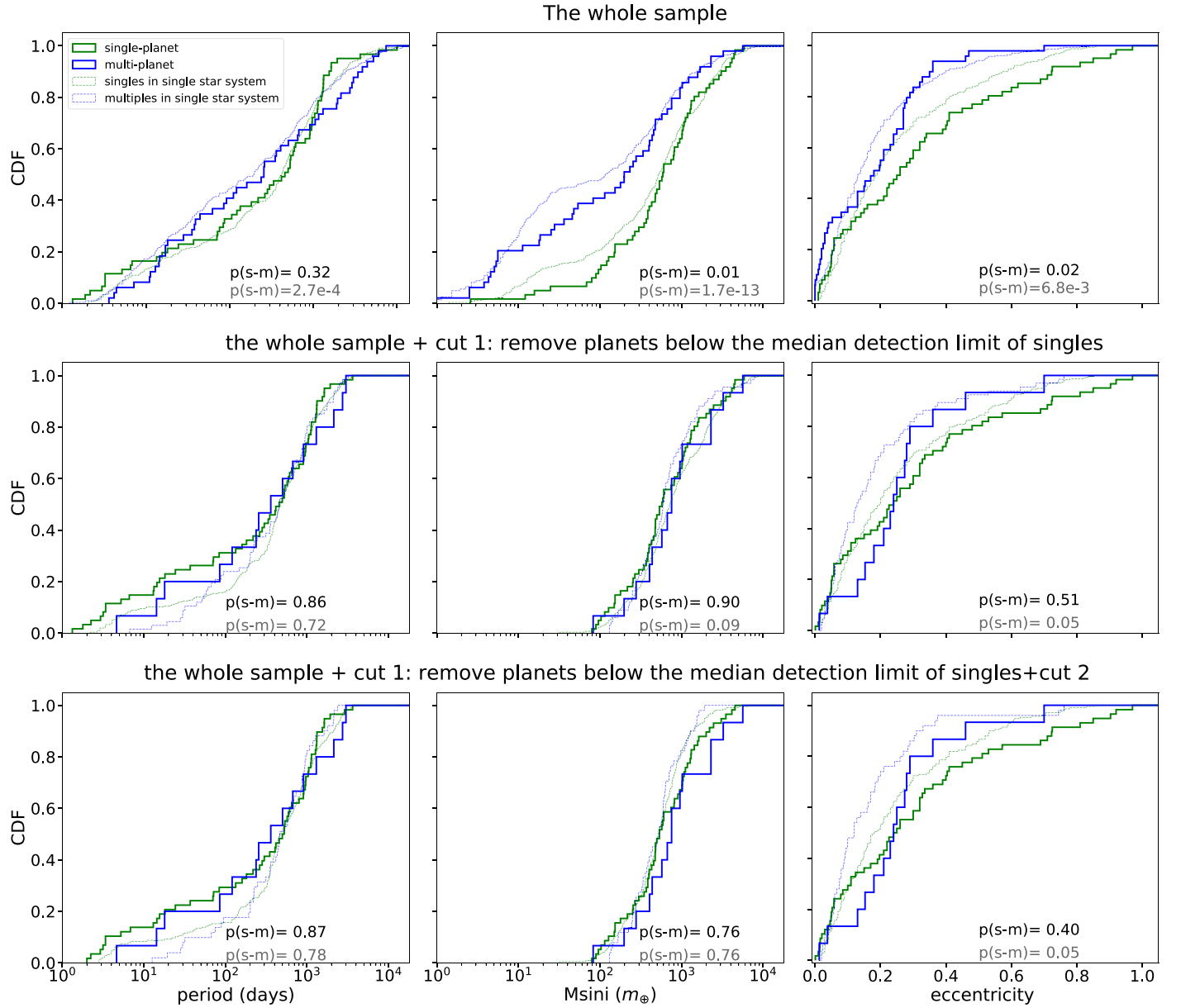


Figure 6. CDFs of planetary orbital periods (first column), minimum masses (M_{sini}, second column), and orbital eccentricities (third column) for single (green) and multiple (blue) planet systems. *Top row:* “raw” sample defined without taking into account potential observational biases. *Middle row:* removing planets below the median detection limit of the singles “cut-1”). *Bottom row:* with both cut-1 and cut-2, i.e., removing systems with RV rms > 100 ms⁻¹. Note that, in the middle and bottom rows, the single- and multiple-planet samples are defined according to the number of planets left after applying the different detection limit cuts (so that, for example, some “multiple systems” of the top row are categorized as “single-planet” systems in the middle row). In each panel, we perform the two sample KS test and print the p value of the test between singles and multiples, i.e., $p(s-m)$. For comparison, in the three right columns, we also plot the results of RV planets in single star systems (light curves) from the NASA Exoplanet Archive and print the corresponding p values (gray numbers).

follows

$$P_{\text{gap}} = \sum_{k=0}^m C_n^k (1-p)^{n-k} p^k, \quad (3)$$

where n is the total points, m is the number of points observed in the gap, C_n^k the combination coefficient and p the nominal probability each point would be located in the gap under the null hypothesis. Throughout this paper, we adopt $p=f=0.22$. Here, there are $n=53$ singles being considered, then the probability that none ($m=0$) is found in the gap by chance is $P_{\text{gap}} = (1-p)^{53} = 2.28 \times 10^{-6}$.

Nevertheless, above simple calculation implicitly assumes that planets in the rectangle are fully detectable (100%

detection completeness) for all single systems, which is not the case in reality. By taking into account the effects of detection completeness of individual systems, the probability that no more than m out of n singles are found in the gap by chance are modified as (marked with a superscript “**”)

$$P_{\text{gap}}^* = \sum_{k=0}^m \sum_{l=1}^k \left(\prod_{i=1}^{n-k} (1-p \times D_{l,i}) \prod_{j=1}^k p \times D_{l,j} \right), \quad (4)$$

where $D_{l,i}$ and $D_{l,j}$ are the coefficients to correct the detectability of individual systems in the l -th case in C_n^k combinations.

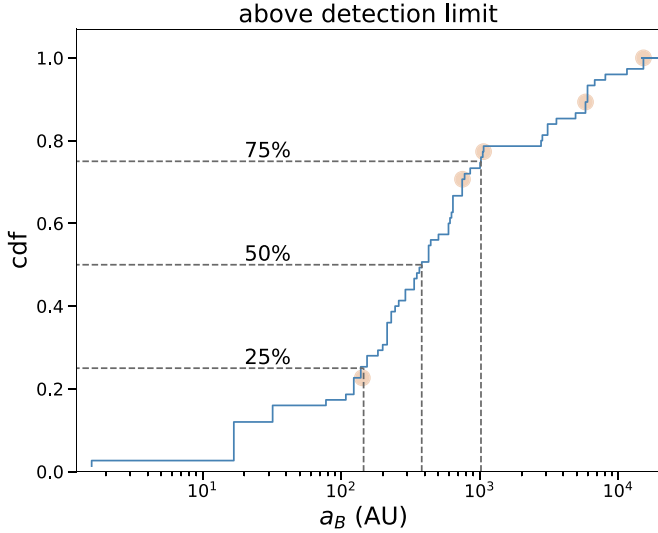


Figure 7. CDF of the orbital semimajor axis a_B of the binaries with planets lying above the detection limit as defined by the green curve in Figure 4. Binaries with planets located in the period-mass gap (the shaded rectangle in Figure 4) are marked with orange circles. For comparison, the 25%, 50% and 75% ($a_B^{25\%} = 155$ au, $a_B^{50\%} = 383$ au, and $a_B^{75\%} = 1020$ au) are also marked.

If we assume that planets, if any, are evenly distributed in the rectangle, then in practice, $D_{l,i} = A_i/A_{tot}$, where A_{tot} is the total area of the rectangle and A_i the area of the part of rectangle that is above the detection limit of individual systems. For example (see Figure 12), HD 41004 B has a poor RV precision ($rms=589.2$ ms^{-1}), and thus zero detectability in the rectangle, i.e., $D_{l,i} = A_i/A_{tot} = 0$; in contrast, HD 41004 A has a good RV precision ($rms=9.69$ ms^{-1}), and 100% detectability in the rectangle, i.e., $D_{l,i} = A_i/A_{tot} = 1$. Using the above Equation (3), the probability that none of singles is found in the gap by chance is $P_{gap}^* = 1.15 \times 10^{-5}$, which is over five times of the p value ($P_{gap} = 2.28 \times 10^{-6}$) if observational biases of individual systems were ignored as mentioned before.

Note, as mentioned before, that there are two criteria for classifying systems as singles and multiples. One is based on the number of planets *actually observed* in the system. The other one is based on the number of remaining planets after applying detection limit cuts. Contrary to the analysis of the distributions of binary separation or planetary orbital parameters (Figures 5 and 6), we here chose the first criteria for defining the two subsamples of single- and multiple-planet systems. The main reason is that the second criterion is statistically too drastic, converting too many multiples (8 out of 23) into singles, resulting in (1) possible contamination of any potential feature (e.g., the gap) that is intrinsic to singles, and (2) too few remaining multiples for statistical study in the 2D period-mass diagram.

Granted, the first criterion also has its own weakness, mainly that some of the observed singles could be misidentified because of potential undetected planets. Nevertheless, the effect of such an observational bias can be relatively straightforward to quantify. Assuming that N_{S-B} observed singles in the left panel of Figure 4 are in reality “hidden” multiples because of the current detection efficiency of multiples just removes N_{S-B} planets around the gap without adding any planet to it in the left panel of Figure 4, the gap itself should basically remain as long as N_{S-B} is not particularly large compared to the total number of singles ($N_S = 61$).

We obtain a rough estimate of N_{S-B} by injecting all the second strongest planet signals from the multiples (right panel of Figure 4) into the RV data of each single (left panel of Figure 4).

We find that, on average, a fraction $\sim 43\%$ of the injected planet signals could be recovered, indicating that $\sim 43\% * N_S = 26$ of the observed singles are likely to be real singles. For the remaining $\sim 57\% * N_S = 35$ uncertain ones, if assuming a single/multiple ratio comparable to that for the certain ones (i.e., $\sim 43\% * N_S / N_M = 26/49$), then 23 of them should be intrinsic multiples, namely $N_{S-B} = 23$. We then estimate how the significance of the gap would be modified. Given $N_{S-B} = 23$ observed singles are reclassified as multiples, the fraction of multiples in the gap reduces to $f = 5/(23 + 23) = 0.10$, and the number of singles reduces to 38. Putting these two reduced numbers into Equations (3) and (4), we obtain $P_{gap} = 0.017$ and $P_{gap}^* = 0.028$ (about 2-sigma confidence). Recalling the previous $P_{gap} = 2.28 \times 10^{-6}$ and $P_{gap}^* = 1.15 \times 10^{-5}$, we therefore conclude that the effect of mixing undetected multiples in observed singles is unlikely to destroy the gap in singles though significantly reduce the statistical confidence.

3.2. Close versus Wide Binaries

3.2.1. Is the Gap More Prominent in Close Binaries?

As there seems to be a link between planetary multiplicity and binary star separation (Section 3.1.1), we here investigate whether the period-mass gap found in single-planet systems (Section 3.1.3) is also dependent on binary separation. As before, we only consider systems with “good detectability”, i.e., systems that lie above the median detection limit of singles (the green line in Figure 4). Even though the gap itself lies in its entirety above the green line, this criterion will change the number of planets outside of the gap that will serve to estimate the statistical significance of the gap. Figure 7 shows the cumulative distribution of the binary separations (a_B) of these systems. There are five binaries which host planets in the gap, and their separations ($a_B = 142, 750, 1065, 5796$ and 15184 au) are marked in Figure 7. Clearly, these gap planets tend to reside in binaries with relatively wide separations, meaning that the gap is preferentially found in close binaries. To further demonstrate this point, we conducted two additional analyses as follows.

In the first analysis, we divided the whole sample into two “close” and “wide” binary subsamples, taking as a boundary the median binary separation for our whole sample $a_B^{50\%} = 383$ au. Figure 8 shows the period-mass diagrams of planets in these two subsamples. As can be seen, in the wide subsample, out of the 37 planets above the detection limit line, there are four lying in the gap, corresponding to a fraction of $\sim 11_{-5}^{+8}\%$ (errorbars reflecting Poisson uncertainties). In contrast, in the close-binaries subsample, there is only 1 planet out of 39 ($\sim 2_{-2}^{+6}\%$) that lies in the gap, and actually near the lower boundary of it.

In the second analysis, we consider a more stringent definition of our close and wide subsamples: a binary is in the close subsample if its a_B is smaller than the 25% of the sample ($a_B^{25\%} = 155$ au), and is in the wide subsample if its a_B is larger than the 75% of the sample ($a_B^{75\%} = 1020$ au). As can be seen in Figure 9, there are now 3 out of 19 ($\sim 15_{-8}^{+15}\%$) planets in wide binaries that reside in the gap. In contrast, the number of gap planets is one (1 out of 20 or $\sim 5_{-4}^{+11}\%$ if considering Poisson uncertainty) for the close-binary subsample.

Both analyses show that the period-mass gap identified for single-planet systems is also preferentially found in close-binaries, even though there is a relatively large Poisson uncertainty given the relatively small size of the subsamples.

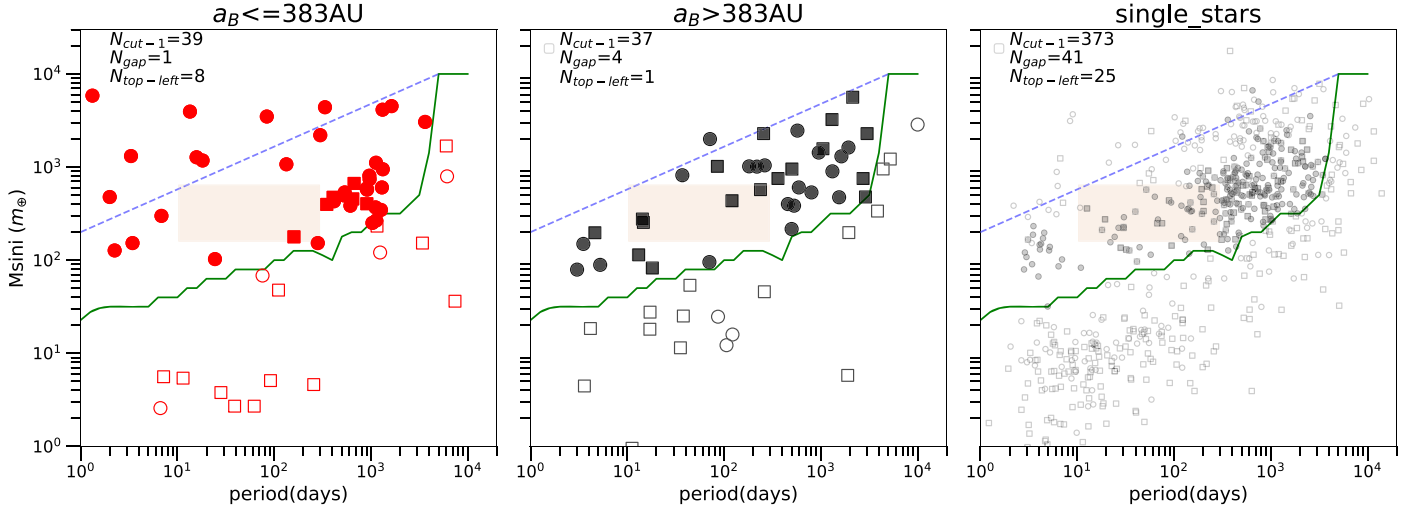


Figure 8. Mass (M_{sini})–period diagrams of RV planets in close binaries (left panel), wide binaries (middle panel) and single star systems (right panel). Here, close/wide binaries are defined as a_B smaller/larger than the 50%, $a_B^{50\%} = 383$ au, of the sample. In each panel, the green curve is the detection limit as defined by the same green curve in Figure 4, the blue dashed line is an empirical fit to the upper envelope of the planet distribution in the wide subsample (panel), and the shaded rectangle markers the same planet desert as identified in Figure 4. Planets above/below the green curve are marked with filled/open circles (single-planet systems) and squares (multiple-planet systems), respectively. In the top left of each panel, we print the number of planets above the detection limit ($N_{\text{cut}-1}$), the number of planets in the gap (N_{gap}), and the number of planets on the top-left side of the blue dashed line ($N_{\text{top-left}}$).

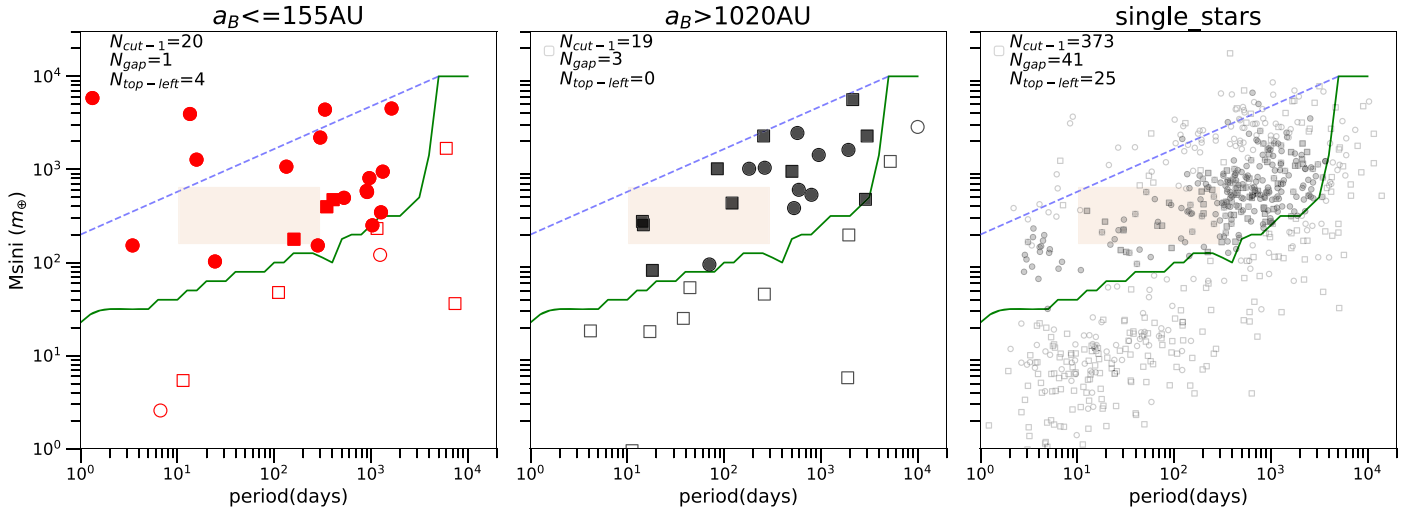


Figure 9. Similar to Figure 8 but here, close/wide binaries are defined as a_B smaller/larger than the 25%/75%, $a_B^{25\%} = 155$ au/ $a_B^{75\%} = 1020$ au, of the sample.

3.2.2. Period, Mass, and Eccentricity Distributions

As can be seen in Figures 10 and 11, there is no statistically significant difference in the 1D distribution of planetary period, mass and eccentricity between systems in close and wide binaries. This result holds for the two different definitions of the close and wide subsamples and regardless whether observational bias cuts are taken into account or not. These results also clearly show up when considering the p values of the KS test between close and wide binaries, which are all greater than 0.5 in Figure 10 and greater than 0.1 in Figure 11.

There is, however, a difference between the 2 subsamples for the 2D distribution of the planetary period-mass diagram, which is different from the already identified gap that prefers close binaries over wide binaries (Section 3.1.3). As can be seen in Figures 8 and 9, there is a rising upper envelope planets in the period-mass diagram for wide binaries that are not found in the close-binary subsample. This upper envelope empirically

follows

$$\log_{10}\left(\frac{M_{\text{sini}}}{m_{\oplus}}\right) = 0.5 \times \log_{10}\left(\frac{\text{period}}{\text{day}}\right) + 2.3. \quad (5)$$

Specifically, as shown in Figure 8, only 1 out of 37 ($\sim 2_{-2}^{+6}\%$) planets (with filled symbols) is above the blue dashed line (an empirical upper envelope) in the wide subsample, while this fraction number is 8/39 ($\sim 21_{-7}^{+10}\%$) in the close subsample. Such a difference is still significant ($0/19 \sim 0_{-0}^{+9}\%$ of wide v.s. $4/20 \sim 20_{-10}^{+16}\%$ of close) in Figure 9, in which close/wide binaries are defined with a more stringent criterion. Note that this result that massive short-period planets are rare in wide binaries cannot be explained by observational bias because these planets are the easiest ones to be detected. The top-left region of the period-mass diagram has indeed the highest detection efficiency as shown in Figure 4.

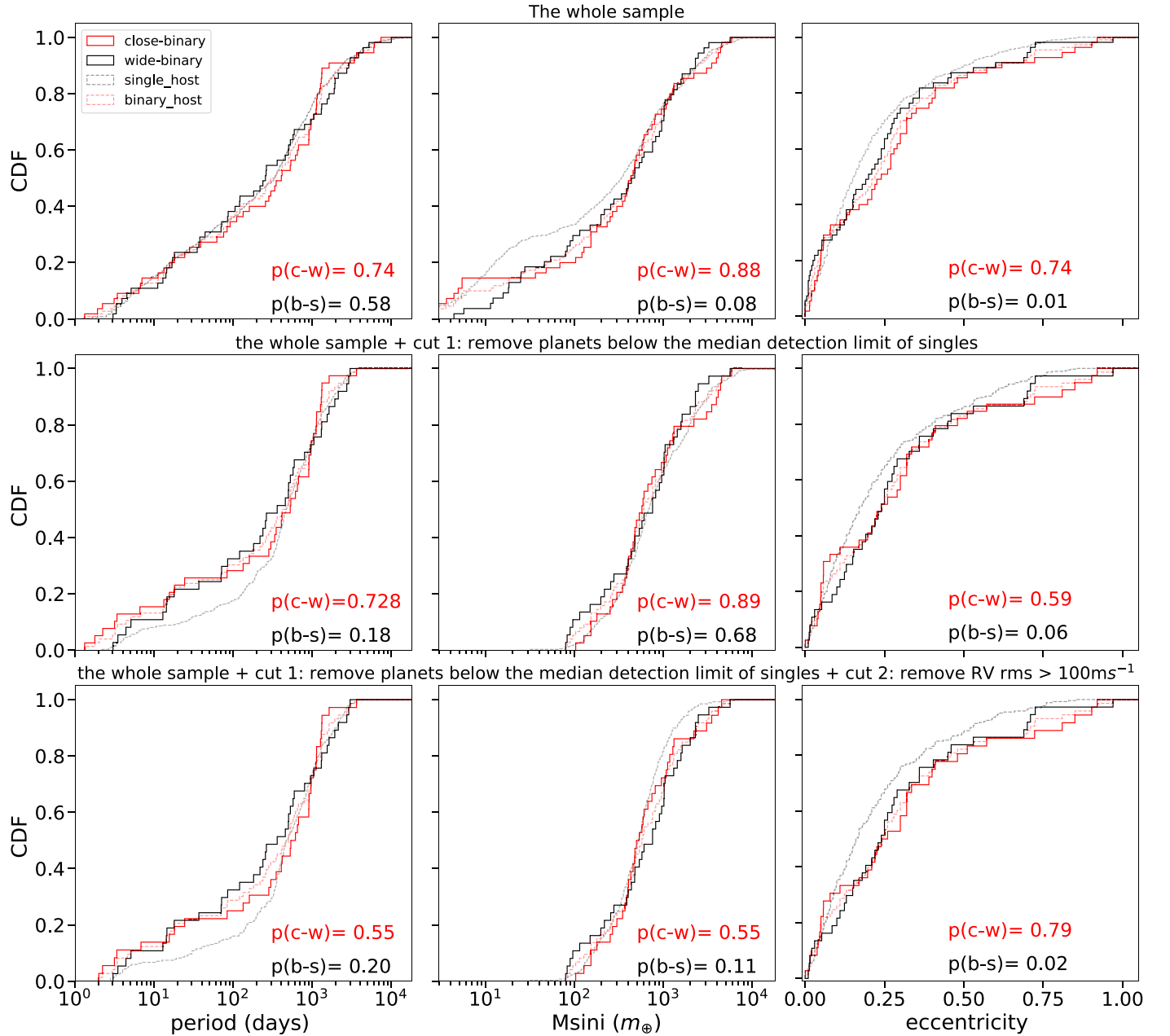


Figure 10. Comparison of planetary orbital periods (first column), minimum masses ($M_{\text{sin}i}$; second column) and orbital eccentricities (third column) between close (red) and wide (gray) binaries in samples without any observational bias cut (top row), with cut-1, i.e., removing planets below the median detection limit of the singles (middle row) and with both cut-1 and cut-2, i.e., removing systems of $\text{RV rms} > 100 \text{ ms}^{-1}$ additionally (bottom row). In each panel, we perform the two sample KS test and print the corresponding p value. Here, close/wide binaries are defined as in Figure 8; namely, a_B smaller/larger than the 50%, $a_B^{50\%} = 383 \text{ au}$ of the sample.

3.3. Single Stars versus Binary Stars

For the sake of comparison, we also performed the above analyses to RV planets in single star systems, which were retrieved from the <https://exoplanetarchive.ipac.caltech.edu> archive.

As can be seen in the three right columns of Figure 6 (dashed curves), the 1D period and mass distributions of planets become indistinguishable between single- and multiple-planet systems after applying the debiasing cut-1 and cut-2, a result similar to the one obtained for binaries. Nevertheless, we note that, contrary to binaries, the 1D distribution in eccentricity remains different between singles and multiples even after debiasing, with a p value of 0.05 (~ 2 -sigma confidence level).

The fraction of planets in the period-mass gap is $41/373 \sim 11^{+2}_{-2}\%$ for single stars (Figures 8 and 9), a value close to the one found for wide binaries ($\sim 11^{+8}_{-5}\%$), strengthening the result found in Section 3.2.1 that the gap is preferentially found in close binaries.

The rising upper envelope in the period-mass diagram identified for planets in wide binaries is also present for planets around single stars (right panels of Figures 8 and 9). There are indeed only 25 out of 373 planets, corresponding to $\sim 6.7^{+1.6}_{-1.3}\%$, above the blue dashed line. This fraction is consistent (within 1-sigma) with the one found for planets in wide binaries ($1/37 \sim 2^{+6}_{-2}\%$), but significantly lower than that of planets in close binaries ($8/39 \sim 21^{+10}_{-7}\%$).

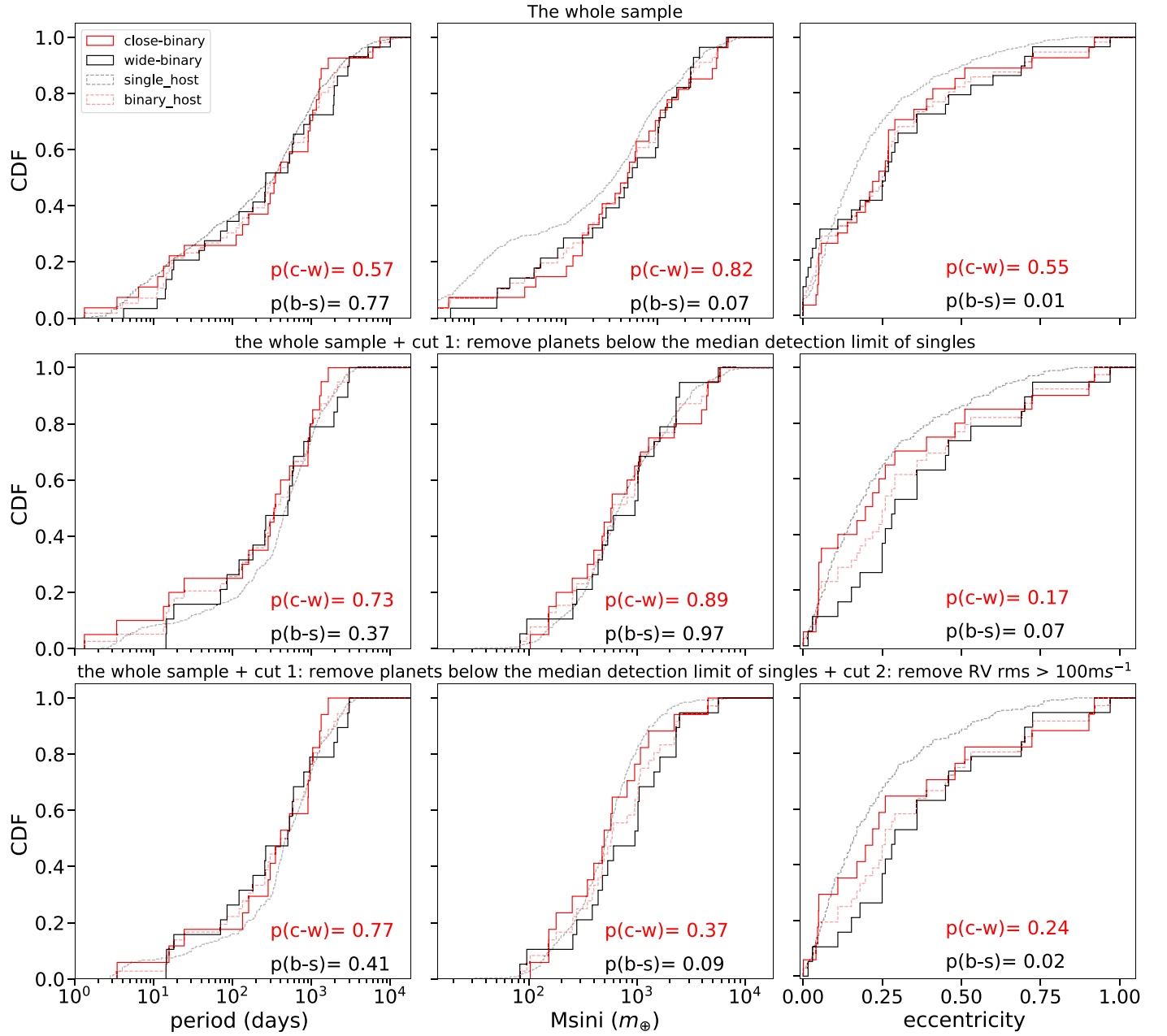


Figure 11. Similar to Figure 10 but here, close/wide binaries are defined as in Figure 9, namely, a_B smaller/larger than the 25%/75%, $a_B^{25\%} = 155 \text{ au}/a_B^{75\%} = 1020 \text{ au}$, of the sample.

As can be seen from the right panels of Figures 10 and 11, planets around single stars tend to have lower eccentricities than planets in binaries. The median planetary eccentricity is 0.16 for single stars, while it is 0.24 in binary star systems, as shown in Figure 10 with a KS test p value of 0.01 – 0.06. The result holds when only considering binaries at the two ends of the separation distribution, i.e., a_B smaller/larger than the 25%/75%, $a_B^{25\%} = 155 \text{ au}/a_B^{75\%} = 1020 \text{ au}$, of the sample. The median eccentricity of planets in binaries increases to 0.26 with KS test p value of 0.01–0.07 as shown in Figure 11. This result indicates that stellar binarity increases the eccentricities of exoplanets to a certain extent, and this effect is not confined to close binaries but still significant even in very wide binaries (e.g., $a_B > 1020 \text{ au}$).

4. Discussions: Implications to Planet Formation and Evolution

Compared to the distribution of planets in wide binaries and single stars, we find that planets in close ($a_B < 100\text{--}300 \text{ au}$) binaries are preferentially associated with three characteristic features: a larger fraction of single planets, a rectangle-shaped gap in the period-mass diagram, and the presence of massive and short-period planets (Figure 6, 7, 8, and 9). We discuss hereafter the possibility that these three features could have the same underlying explanation.

Close binaries with a separation of less than a few hundred au are indeed capable of enhancing planet migration either in the early gas-disk phase (Kley 2000) or in the later disk-free phase, e.g., via secular planetary interactions (Fabrycky & Tremaine 2007;

Wu & Lithwick 2011), causing planets to horizontally cross the upper rising envelope in the period-mass diagram identified in Figures 8 and 9. In addition, these planets could accrete additional material (e.g., gas, planetesimals or other smaller planets) during their enhanced migration, and could thus grow into much more massive objects (Kley 2000; Zhang et al. 2018). As a consequence, these planets would populate the initially depleted upper-left part of the period-mass diagram. These enhanced migrations could also lead to an enhanced level of planetary mergers and/or ejections, explaining the lower level of planet multiplicity in close binaries, as well as their larger eccentricities compared to those in single star systems. As for the rectangle-shaped period-mass gap, it could be the birth place of those line-crossing planets or planets that were accreted/ejected by them.

As for planets in wide binaries, they share a number of common properties with those around single stars, namely a similar fraction (about 10%) of planets in the period-mass gap combined to the presence of a similar upper envelope in the same period-mass diagram (Figures 8 and 9). The rising upper envelope is consistent with the positive period-mass correlation reported by previous studies (Zucker & Mazeh 2002; Jiang et al. 2007) and with the predictions of current synthetic models (Ida & Lin 2004; Mordasini et al. 2009) of planet formation around single stars. The common upper envelope between planets in wide binaries and single stars thus may suggest that planet formation and evolution in wide binaries ($a_B > 100\text{--}300$ au) is mostly similar to that in single star systems.

Nevertheless, planets in wide binaries differ significantly from planets in single star systems in terms of their eccentricity distribution (Figures 6, 8, and 9). For single stars, multiple-planet systems have significantly lower eccentricities than single-planet systems (the so called ‘‘eccentricity dichotomy’’), a fact that has been well established for the Kepler planet sample (dominated by super-Earths and sub-Neptunes; Xie et al. 2016; Van Eylen et al. 2019; Mills et al. 2019) and is also seen in RV planets (dominated by Jovian planets) as shown here in Figure 6. In contrast, there is no difference (once the data is debiased) between the eccentricity distributions of single- and multiple-planet systems for both our close- and wide-binary subsamples. In addition, these eccentricities are, on average, larger than the eccentricities of planets around single stars (Figures 6, 10, and 11). This may imply that there are some additional mechanisms at play to pump up planet eccentricities in wide binaries and bring them to the same level as for those in closer binaries. Potential eccentricity drivers could be the Kozai mechanism (Kozai 1962) and Galactic tides (Kaib et al. 2013). Future studies, both with larger observational sample and with dedicated numerical models are needed to further unveil the picture of planet formation and evolution in binary star systems.

5. Summary and Conclusion

In this paper, we compile a sample (Table 1) of all S-type planetary systems detected by the RV method in order to derive statistical characteristics as a function of key properties of these systems, such as the number of detected planets or binary separation. Available RV observation data allow us to quantify

the planet detection efficiencies of individual systems, enabling us to correct for potential observational bias, notably between the single-planet and multiple-planet subsamples. We then perform a statistical investigation from two different perspectives.

First, we consider the whole sample of S-type systems and look for statistical properties that depend on planetary multiplicity. We find that, for the whole S-type sample, single and multiple-planet systems have similar 1D distributions of both planetary masses and periods. However, they differ significantly in terms of their 2D period-mass distribution. The main difference is an absence of planets in a rectangle-shaped gap (for planetary periods between 10 and 300 days M_{sin} between 160 and 640 $m_{\oplus} \sim 0.5\text{--}2.0 m_J$) that is present in single-planet systems but not in multiples. The gap is statistically significant with a P value of $P_{\text{gap}} = 2.28 \times 10^{-6}$ or $P_{\text{gap}}^* = 1.15 \times 10^{-5}$ depending if observational bias is ignored (Equation (3)) or considered (Equation (4)).

Second, we look for planet properties that depend on binary separation (a_B) and divide our sample into a close and wide binary subsamples. We find that multiple-planet systems are preferentially found in wide binaries, which is consistent with the previous study by Roell et al. (2012). We also find that the gap in the planetary period-mass diagram of single-planet systems depends on binary separation: it is more prominent in close binaries with a_B less than a few hundreds au than in wide binaries or single stars. Furthermore, we find that, in the same planetary period-mass distribution, there is a population of massive and short-period planets in close-binaries that is not found in wide binaries and single stars, for which the period-mass distribution is limited by a rising upper envelope. Last but not least, we find that the average eccentricity of planets in binaries (both close and wide) is higher than in single stars.

We suggest that enhanced planetary migration, collision and/or ejection in close binaries could explain most of the statistical properties we have identified, but the detailed study of this effect exceeds the scope of the present paper.

In summary, our study emphasizes the importance of taking into account individual detection efficiencies when searching for statistical patterns for planets in binaries. The results of this paper suggest that binary stars could play a crucial role in shaping the architecture of planetary systems, i.e., reducing planetary multiplicity, exciting planetary eccentricity and modifying the planetary period-mass diagram.

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Appendix

RV-signal Injection-recovery Tests of the Whole Sample

In Figures 12–15, we present the results of RV-signal injection-recovery tests for all the 80 systems (110 planets) studied in this paper.

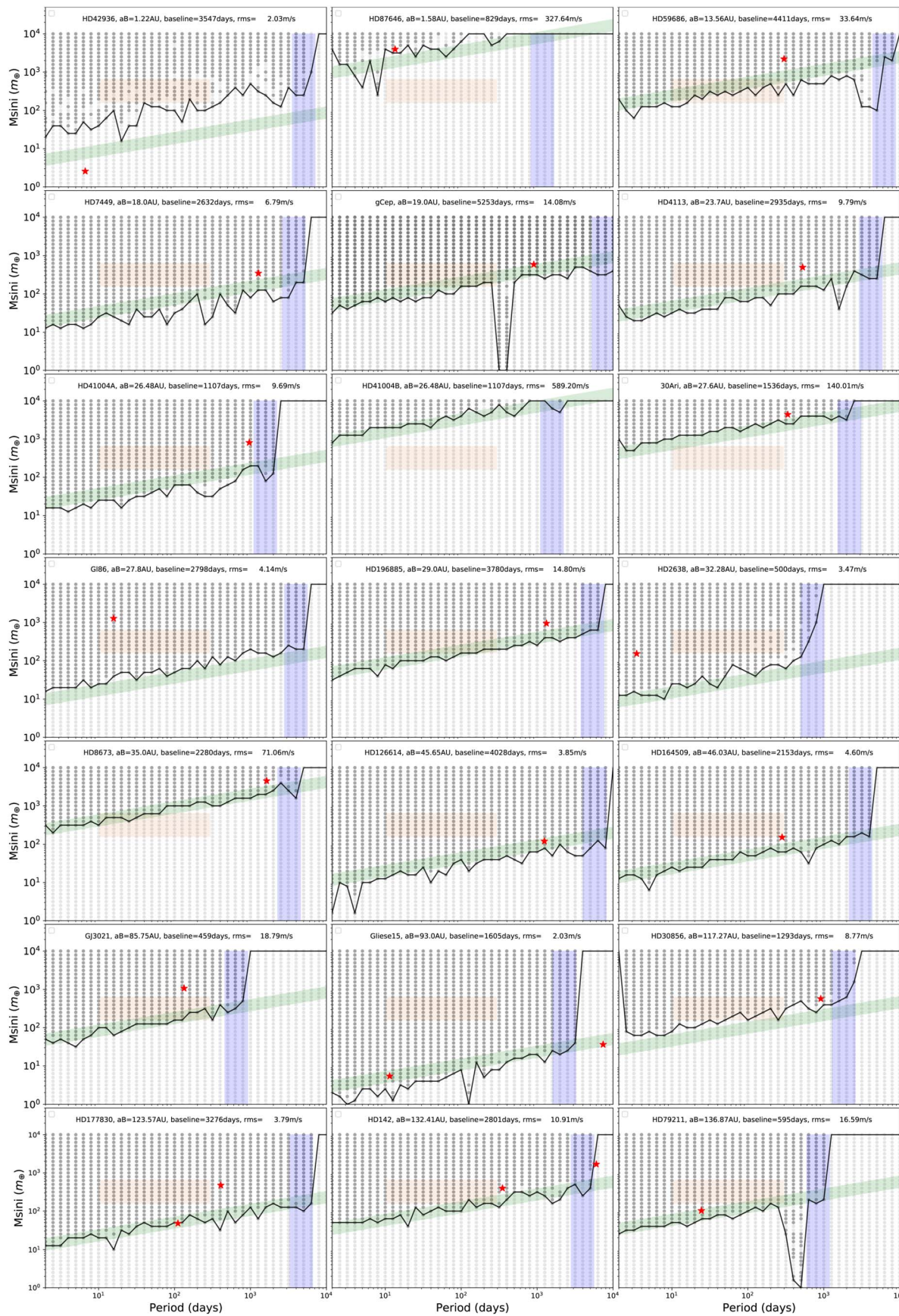


Figure 12. RV-signal injection-recovery tests of 21 binary systems with $a_B \leq 137$ au. The red star symbols indicate the location of observed planets. The filled dark-gray circles (light squares) show where the hypothetical planets can (cannot) be recovered from the simulated RVs. The green shaded regions mark where $RV_{\text{signal}} = (1 - 2) \times rms$. The purple shaded regions mark where $period = (1 - 2) \times \text{the time baseline}$ of the simulated RV data. The black curve mark the detect limit of the binary system.

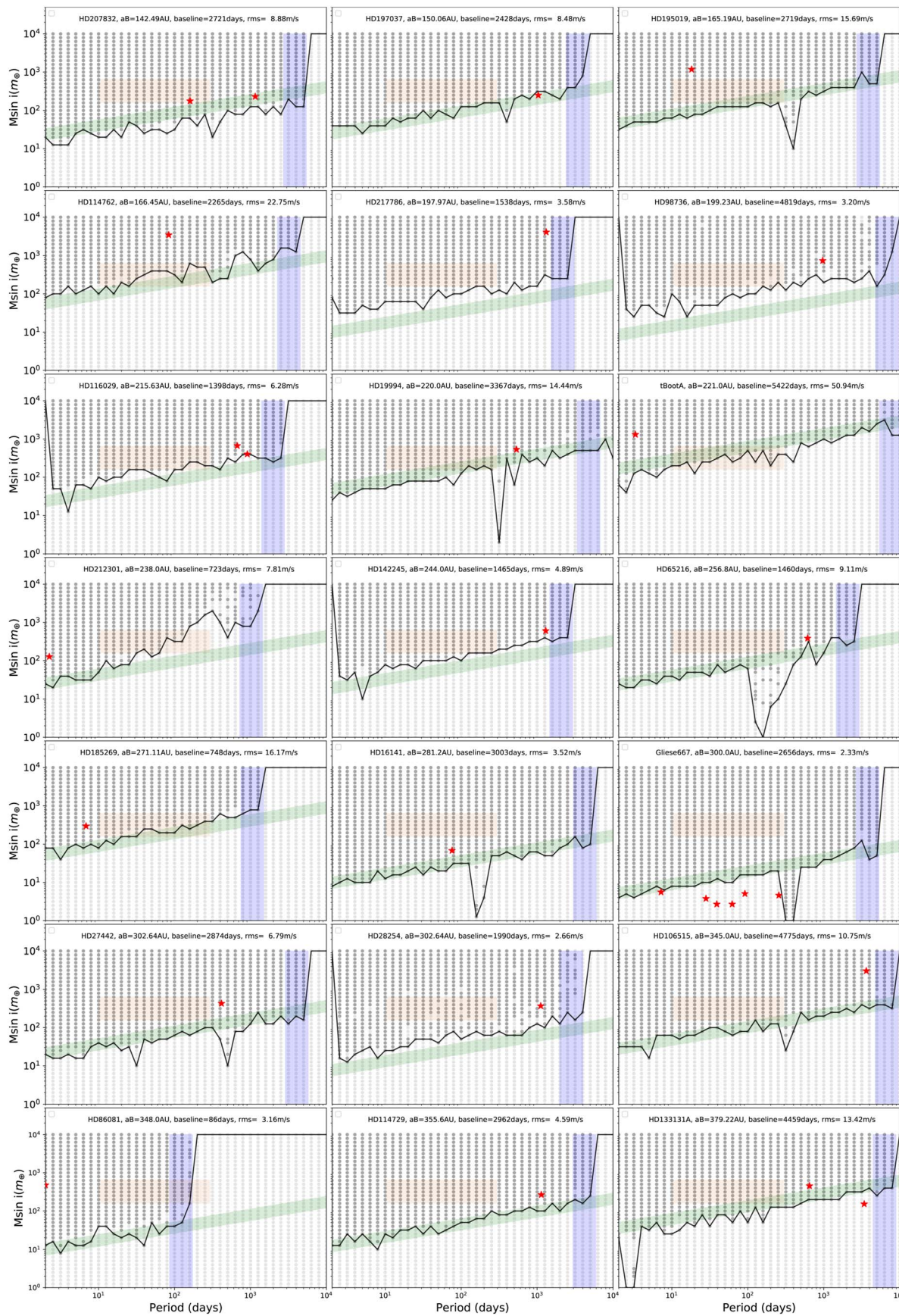


Figure 13. Similar to Figure 12 but showing the RV-signal injection-recovery tests of 21 binary systems with $142 \text{ au} < a_B \leq 380 \text{ au}$.

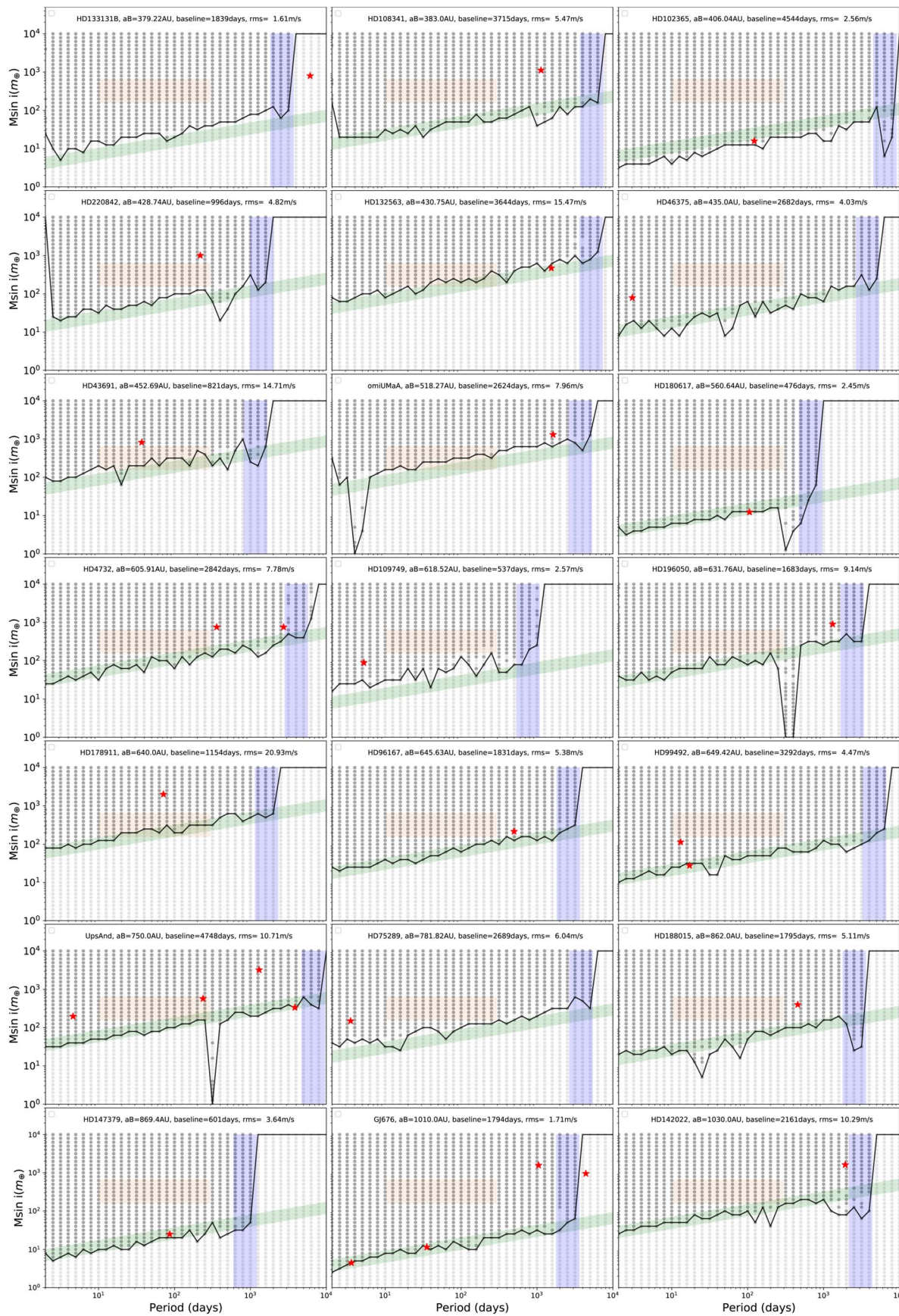


Figure 14. Similar to Figure 12 but showing the RV-signal injection-recovery tests of 21 binary systems with $380 \text{ au} < a_B \leq 1030 \text{ au}$. The caption is similar to Figure 12.

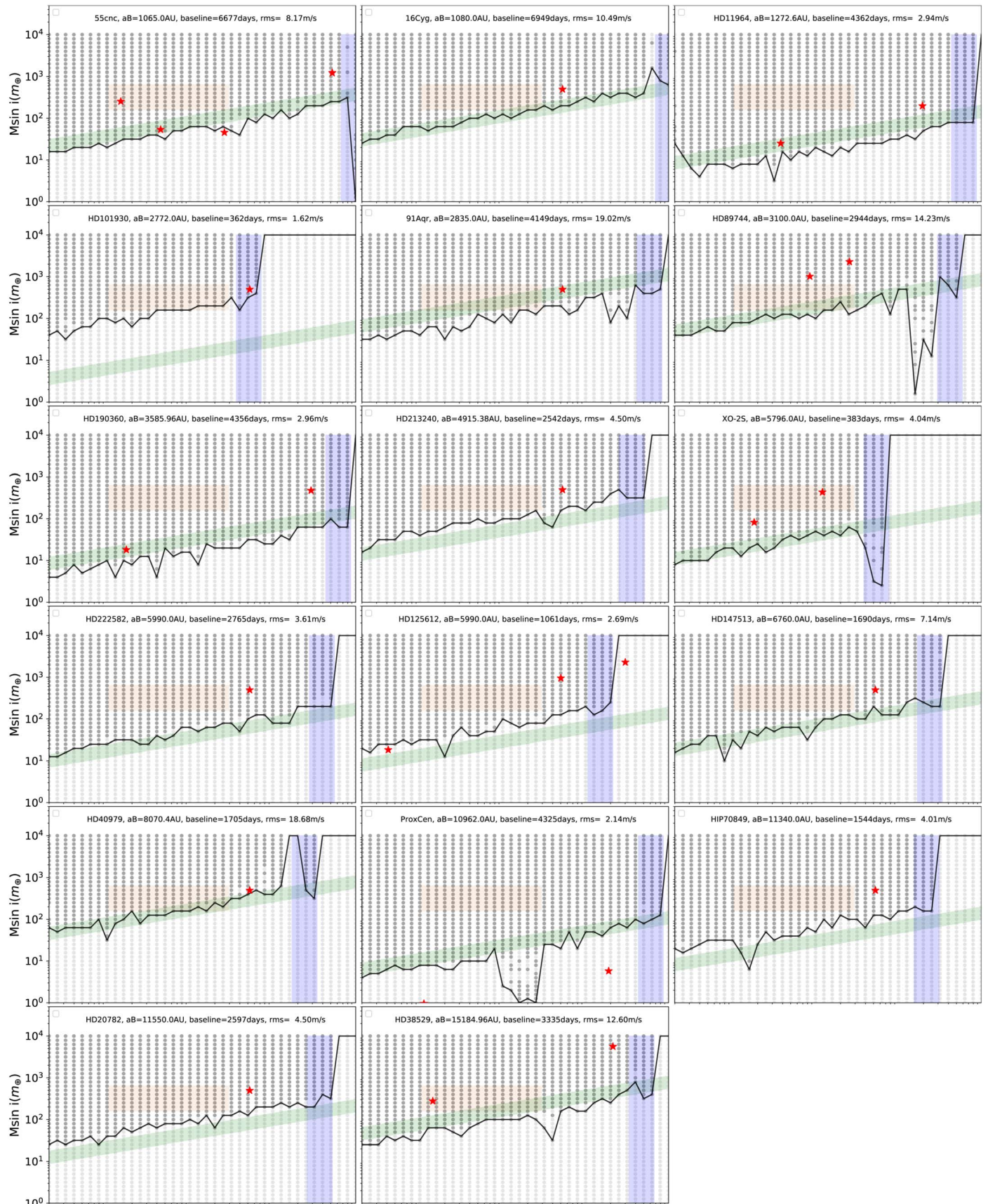





Figure 15. Similar to Figure 12 but showing the RV-signal injection-recovery tests of 17 binary systems with $a_B > 1030$ au. The caption is similar to Figure 12.

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